Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 000000000000

On Sequential Monte Carlo (SMC) strategies for Target Distributions

M. Rousset 1,2

F. Cérou, A. Guyader, B. Delyon, T. Lelièvre, G. Stoltz, C.E. Bréhier, L. Goudenège, P. Héas. (PhDs: F. Ernoult, K. Tit).

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Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
Aim of the	talk			

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Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
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• 'Target probability distribution': defined as a density w.r.t to a simulable distribution, density given up to a normalizing constant. E.g.: posterior or conditional distribution, Gibbs probability.

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- 'Target probability distribution': defined as a density w.r.t to a simulable distribution, density given up to a normalizing constant. E.g.: posterior or conditional distribution, Gibbs probability.
- SMC = particle methods= Importance splitting. As "opposed" to MCMC methods. Start with a sample of *N* replicas ('particles'). Algorithms output: sample of *N* particles (approx. indep.) with distribution the 'target' + estimator of normalisation.

Context and Algo ●○○○○○○○○○○○○○	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
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- SMC = particle methods= Importance splitting. As "opposed" to MCMC methods. Start with a sample of *N* replicas ('particles'). Algorithms output: sample of *N* particles (approx. indep.) with distribution the 'target' + estimator of normalisation.
- Aim of the talk: Overview on variants and adaptivity + app. to neutrons (cf. talk of Lelièvre and Del Moral).

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000

E.g.: Rare event problem

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Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
E.g.: Rare	event problem			

• $\pi(dx)$ a reference probability on S that can be simulated.



Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
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• $\operatorname{score}: \mathbb{R}^d \to \mathbb{R}$ a given computable function.

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
E a · Para	avent problem			

- $\pi(dx)$ a reference probability on S that can be simulated.
- score : $\mathbb{R}^d \to \mathbb{R}$ a given computable function.
- Assume $\pi(\{\text{score} > 0\}) = 1$. Problem for given *s*:

 $\begin{array}{ll} \mbox{Estimate} & p_s := \pi(\{\mbox{score} > s\}) \ll 1 \\ \mbox{Simulate according to } \underline{\ 'target'} & \eta_s(dx) := \pi(dx|\mbox{score}(x) > s). \end{array}$

Idea (For high dimensions / low temperature)

Estimate/Simulate "smoothly" and sequentially the path

$$s\mapsto(p_s,\eta_s),\quad s\in[0,1].$$

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
Generalizat	ion			

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Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
Generalizat	ion			

• $\frac{1}{z_0}e^{-V_0(0)}\pi(dx)$ a reference probability on $S = \mathbb{R}^d$ that can be simulated, e.g. $z_0 = 1$.

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- $\frac{1}{z_0}e^{-V_0(0)}\pi(dx)$ a reference probability on $S = \mathbb{R}^d$ that can be simulated, e.g. $z_0 = 1$.
- (s,x) → V_s(x) : ℝ × ℝ^d × → ℝ a given computable function (called potential). (Optional: ∇_xV_s(x) is available).

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Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000	
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- $\frac{1}{z_0}e^{-V_0(0)}\pi(dx)$ a reference probability on $S = \mathbb{R}^d$ that can be simulated, e.g. $z_0 = 1$.
- (s,x) → V_s(x) : ℝ × ℝ^d × → ℝ a given computable function (called potential). (Optional: ∇_xV_s(x) is available).
- Problem, for s := 1:

Estimate the normalization: $z_s := \pi(e^{-V_s()})$ Simulate according to 'target': $\eta_s(dx) := \frac{1}{z_s}e^{-V_s(x)}\pi(dx)$.

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000

- $\frac{1}{z_0}e^{-V_0(0)}\pi(dx)$ a reference probability on $S = \mathbb{R}^d$ that can be simulated, e.g. $z_0 = 1$.
- (s,x) → V_s(x) : ℝ × ℝ^d × → ℝ a given computable function (called potential). (Optional: ∇_xV_s(x) is available).
- Problem, for s := 1:
- Previous rare event model is particular case for:

$$V_s(x) = \begin{cases} +\infty & \text{if score}(x) \leqslant s \\ 0 & \text{if score}(x) > s \end{cases}$$

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Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000

¹Lelièvre-Stoltz-Rousset, Langevin dynamics with constraints and computation of free energy differences, 2012

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• $\frac{1}{z_0}e^{-V(x,0)}\pi_0(dx)$ a target probability on $S = \mathbb{R}^d$ that can be simulated. $z_0 = 1$.

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Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
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- $\frac{1}{z_0} e^{-V(x,0)} \pi_0(dx)$ a target probability on $S = \mathbb{R}^d$ that can be simulated. $z_0 = 1$.
- Target : $e^{-V_s} d\pi_s/z_s$.

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- $\frac{1}{z_0} e^{-V(x,0)} \pi_0(dx)$ a target probability on $S = \mathbb{R}^d$ that can be simulated. $z_0 = 1$.
- Target : $e^{-V_s} d\pi_s/z_s$.
- $s \mapsto \pi_s$ a path of mutually singular non-negative reference measures and a family of computable maps $i_{s,s'} : \mathbb{R}^d \to \mathbb{R}^d$ with $s, s' \in \mathbb{R}$ such that:

$$\pi_{s'} = i_{s,s'}[\pi_s]$$
 (push-forward)

Example

 $\pi_s := 2d' < 2d$ -dimensional phase-space volume of a parametric family of co-tangent spaces $s \mapsto T^*\Sigma_s \subset \mathbb{R}^{2d}$. $i_{s,s'}$ is a simulable symplectic projection.

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Adaptivity and Mutations

Bias and consistency

Indexing using selection 0000

Remark: High Dimensional/Low Temperature Applications

Motivated by pbs where:

Relative Entropy(target | educated proposal) $\rightarrow +\infty$.

(that is Importance Sampling fails, see also Chatterjee, Diaconis AAP 2018) .

Adaptivity and Mutations

Bias and consistency

Indexing using selection

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• Sampling w.r.t. low temperature Gibbs distribution. Tempering: $\pi_s \propto e^{-sU(x)}\pi(dx)$.

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- Sampling w.r.t. low temperature Gibbs distribution. Tempering: $\pi_s \propto e^{-sU(x)}\pi(dx)$.
- High dimensional Bayesian statistics: $\pi = \text{prior distribution on model}(s)$. $-V(s, x) = (\text{smoothed}) \text{ log-likelihood from } s \times n_{\text{obs}} \text{ datas.}$

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- Rare event.

Adaptivity and Mutations

Bias and consistency

Indexing using selection

AMS 000000000000

Branching Neutron Transport Application



Adaptivity and Mutations

Bias and consistency

Indexing using selection 0000

AMS 000000000000

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Adaptivity and Mutations

Bias and consistency

Indexing using selection

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Rare Event formalism:



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Bias and consistency

Indexing using selection

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Branching Neutron Transport Application

Rare Event formalism:

• Exemple of choice of score function, e_0 given critical energy:

$$\operatorname{score}(X) := \max_{t \ge 0} \zeta(X_t)$$
$$= \max_{t \ge 0} \inf \left\{ s \in \mathbb{R} \mid t \text{ for a transformation } t \text{ for a transformati$$

total energy of neutrons at time t in $\{\xi \ge s\} \le e_0$

Adaptivity and Mutations

Bias and consistency

Indexing using selection

Branching Neutron Transport Application

Rare Event formalism:

• Exemple of choice of score function, e_0 given critical energy:

$$score(X) := \max_{t \ge 0} \zeta(X_t)$$
$$= \max_{t \ge 0} \inf \left\{ s \in \mathbb{R} \mid total energy of neutrons at time t in \{ \xi \ge s \} \leqslant e_0 \right\}$$

• Corresponding choice of rare event flow:

 $z_s = p_s := \pi(\{ \text{score} \ge s \}), \qquad \eta_s := \pi(\mid \text{score} \ge s)$

Adaptivity and Mutations

Bias and consistency

Indexing using selection 0000

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Branching Neutron Transport Application



Adaptivity and Mutations

Bias and consistency

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Adaptivity and Mutations

Bias and consistency

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Branching Neutron Transport Application

• Define the stopping time for the natural filtration of the branching neutron transport process X

$$\tau_s(X) := \inf \{t \ge 0 \mid \zeta(X_t) = s\}$$

Note that

$$\left\{\operatorname{score}(X) < s\right\} = \left\{\tau_s(X) = +\infty\right\}.$$

Adaptivity and Mutations

Bias and consistency

Indexing using selection

AMS 00000000000

Branching Neutron Transport Application

• Define the stopping time for the natural filtration of the branching neutron transport process *X*

$$\tau_s(X) := \inf \left\{ t \ge 0 \mid \zeta(X_t) = s \right\}$$

Note that

$$\left\{\operatorname{score}(X) < s\right\} = \left\{\tau_s(X) = +\infty\right\}.$$

 Using the strong Markov property, one can simulate new neutrons starting from X_{τs(X)}. This yields a Markov kernel M_s(x, dx') leaving π conditioned by {score ≥ s} invariant.

$$\int M_{\mathfrak{s}}(x,dx')\eta_{\mathfrak{s}}(dx)=\eta_{\mathfrak{s}}(dx').$$

Adaptivity and Mutations

Bias and consistency

Indexing using selection 0000

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Branching Neutron Transport Application



Adaptivity and Mutations

Bias and consistency

Indexing using selection

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Sequential Monte-Carlo a.k.a. Importance Splitting

Define: $0 = s_{(0)} < \ldots < s_{(i_{\max})} = 1$ a given, finite ladder of scores.

 $X_{s_{(i)}}^n$ state of replica *n* at iteration *i*.

General Form of the Algorithm with Weighted Replicas:

Adaptivity and Mutations

Bias and consistency

Indexing using selection

AMS 000000000000

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• (0) Simulate N independent replicas according to η_0 .

Iterate on $i = 1 \dots i_{max}$:

Adaptivity and Mutations

Bias and consistency

Indexing using selection

AMS 000000000000

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• (0) Simulate N independent replicas according to η_0 .

Iterate on $i = 1 \dots i_{max}$:

• (i) Weights: update the 'importance weight' of each replica $n \in (1, N)$ by $e^{-V_{s_{(i)}}(X_{s_{(i-1)}}^n)+V_{s_{(i-1)}}(X_{s_{(i-1)}}^n)}$ (target: $e^{-V_{s_{(i)}}}\pi$).

Adaptivity and Mutations

Bias and consistency

Indexing using selection

AMS 000000000000

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- (*ii*) Selection (optional) kill and/or split replicas and update weights. E.g.: triggered if weights are too degenerate.
Adaptivity and Mutations

Bias and consistency

Indexing using selection

AMS 00000000000

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- (*ii*) Selection (optional) kill and/or split replicas and update weights. E.g.: triggered if weights are too degenerate.
- (iii) Mutation (optional): modify ('mutate') (all or some or none) replicas with Markov Chain Monte Carlo transition M_{s(i)}(x, dx') that leaves invariant the target η_{s(i)}(dx).

Adaptivity and Mutations

Bias and consistency

Indexing using selection 0000

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Sequential Monte-Carlo a.k.a. Importance Splitting

Estimators:

Adaptivity and Mutations

Bias and consistency

Indexing using selection 0000

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Sequential Monte-Carlo a.k.a. Importance Splitting

Estimators:

• Target measures $\eta_s = \frac{1}{z_s} e^{-V(x,s)} \pi(dx)$ are estimated by weighted empirical measures with normalization

$$\eta^N_{s_{(i)}} := \sum_{n=1}^N \operatorname{Weight}_{s_{(i)}}^n \delta_{X^n_{s_{(i)}}} / \sum_{n=1}^N \operatorname{Weight}_{s_{(i)}}^n.$$

Adaptivity and Mutations

Bias and consistency

Indexing using selection 0000

AMS 00000000000

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Normalizations are estimated by the average weights over replicas

$$z^N_{s_{(i)}} := \frac{1}{N} \sum_{n=1}^{N} \operatorname{Weight}^n_{s_{(i)}}$$

Context and Algo ○○○○○○○○○○○○	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000

Definition

A selection or re-sampling scheme draw branching numbers $B_n \in \mathbb{N}$ splitting or killing each replica $n = 1 \dots N$. New weights are defined accordingly:

 $\mathbb{E}[\operatorname{Weight}_{new}^{n}B^{n} \mid \operatorname{Weight}] = \operatorname{Weight}^{n}$

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so that the Effective Sample Size (e.g. Renyi entropy between weighted and unweighted empirical distrib.) increases.

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000

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$\mathbb{E}[\operatorname{Weight}_{new}^{n}B^{n} \mid \operatorname{Weight}] = \operatorname{Weight}^{n}$

so that the Effective Sample Size (e.g. Renyi entropy between weighted and unweighted empirical distrib.) increases.

• $B^n \ge 1$: selection of splitting type.

Context and Algo ○○○○○○○○○○○○	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000

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- $B^n \ge 1$: selection of splitting type.
- $B^n \leq 1$: selection of killing type.

Context and Algo ○○○○○○○○○○○○	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000

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so that the Effective Sample Size (e.g. Renyi entropy between weighted and unweighted empirical distrib.) increases.

- $B^n \ge 1$: selection of splitting type.
- $B^n \leq 1$: selection of killing type.
- $B^n \ge 1$ and $\mathbb{E}(B_n)$ is independent on n: neutral bearing.

Context and Algo ○○○○○○○○○○○○	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
Basic Refs				

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Papers:

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• Del Moral Doucet Jasra *Sequential Monte Carlo samplers* 2006.

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• F Cérou, A Guyader, *Adaptive Multilevel Splitting for rare event analysis*, 2007.

Context and Algo ○○○○○○○○○○○○	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
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- F Cérou, A Guyader, *Adaptive Multilevel Splitting for rare event analysis*, 2007.
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Context and Algo ○○○○○○○○○○○○	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
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Context and Algo ○○○○○○○○○○○○	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
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Books

• Liu Monte Carlo Strategies

Context and Algo ○○○○○○○○○○○○	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
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- Liu Monte Carlo Strategies
- Chopin Introduction To Sequential Monte Carlo

Context and Algo ○○○○○○○○○○○○	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
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- Freddy Bouchet and al..

Books

- Liu Monte Carlo Strategies
- Chopin Introduction To Sequential Monte Carlo
- Doucet, Freitas, Gordon Sequential Monte Carlo in Practice

Context and Algo ○○○○○○○○○○○○	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
Basic Refs				

- Del Moral Doucet Jasra *Sequential Monte Carlo samplers* 2006.
- A Beskos, A Jasra, N Kantas, A Thiery On the convergence of adaptive sequential Monte Carlo methods 2016
- F Cérou, P Del Moral, T Furon, A Guyader *Sequential Monte Carlo for rare event estimation* 2012
- F Cérou, A Guyader, *Adaptive Multilevel Splitting for rare event analysis*, 2007.
- In Phys.: 'Jarzynski equality'
- Freddy Bouchet and al..

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• Del Moral *Feynman-Kac formula*

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
What is 'ac	laptivity' ?			

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Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
What is 'ac	antivity' ?			

• A 'non-adaptive' SMC/Importance Splitting algorithm consist of: i) preset ladder of scores $0 = s_{(0)} < \ldots < s_{(i_{\max})} = 1$, ii) preset choice of mutations M_s leaving target η_s invariant.

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
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- A 'non-adaptive' SMC/Importance Splitting algorithm consist of: i) preset ladder of scores $0 = s_{(0)} < \ldots < s_{(i_{\max})} = 1$, ii) preset choice of mutations M_s leaving target η_s invariant.
- Many 'adaptive' variants (e.g. Adaptive Multilevel Splitting, see after) are presented as follows: the choice of the scores is random, adaptive.

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
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- Many 'adaptive' variants (e.g. Adaptive Multilevel Splitting, see after) are presented as follows: the choice of the scores is random, adaptive.
- In this talk I propose a different 'mindset':

Idea

Scores are always deterministic -> continuous ladder.

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Interpret 'Adaptivity' = 'Mutations Only If Selection'.
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Nothing happens for most scores ('weighting = waiting') because of the triggering of mutations by selection.

Context and Algo	Adaptivity and Mutations ○●○○○	Bias and consistency	Indexing using selection	AMS 00000000000
Triggered N	Autations			

Triggered Mutations	

• Preset Mutations: *M_s* is preset, applied to all replicas at each score → non-adaptive, 'Feynman-Kac-Del Moral structure'.

Context and Algo	Adaptivity and Mutations ○●○○○	Bias and consistency	Indexing using selection	AMS 00000000000
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 Preset Mutations: M_s is preset, applied to all replicas at each score → non-adaptive, 'Feynman-Kac-Del Moral structure'.

• Mutations-If-Selection: A mutation kernel *M_s* is triggered only when selection step is triggered.

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
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- Preset Mutations: M_s is preset, applied to all replicas at each score → non-adaptive, 'Feynman-Kac-Del Moral structure'.
- Mutations-If-Selection: A mutation kernel *M_s* is triggered only when selection step is triggered.
- Mutations-On-Child: The mutation kernel M_s applied only to children when selction is such that the sample of children \simeq target.

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000

Example (Mutations-If-Selection)



Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
Triggered N	Autations			

Example (Mutations-If-Selection)

• Compute an Effective Sample Size (ESS) of the weights at each score/iteration.

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 00000000000
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• If ESS is above a treshhold: trigger selection.

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 000000000000
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Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 000000000000
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• Special case of Mutations-If-Selection.

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 000000000000
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- Compute an Effective Sample Size (ESS) of the weights at each score/iteration.
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Example (Mutations-On-Child)

- Special case of Mutations-If-Selection.
- Resampling/selection is such that for splitting:

$$\operatorname{Law}(Child) = \frac{1}{N} \sum_{n=1}^{N} W^n \delta_{X^n} \simeq \operatorname{target}.$$

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 000000000000
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- Special case of Mutations-If-Selection.
- Resampling/selection is such that for splitting:

$$\operatorname{Law}(\mathit{Child}) = \frac{1}{N} \sum_{n=1}^{N} W^n \delta_{X^n} \simeq \operatorname{target}.$$

• Triggered mutations are applied on children ONLY.

Adaptivity and Mutations

Bias and consistency

Indexing using selection

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Adaptive/Triggered Mutation variant

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Adaptivity and Mutations

Bias and consistency

Indexing using selection

AMS 00000000000

Adaptive/Triggered Mutation variant

Remarks

 The goal of triggered Mutations (If-Selection, On-Child) is to save computational power by avoiding mutations (hence evaluation of V or ∇V) if simple weighting is sufficient.
Adaptivity and Mutations

Bias and consistency

Indexing using selection

AMS 00000000000

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- Consistency of Adaptive mutations: large sample $N \to +\infty$.

Adaptivity and Mutations

Bias and consistency

Indexing using selection

AMS 00000000000

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Adaptivity and Mutations

Bias and consistency

Indexing using selection

AMS 00000000000

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- Well-known rare event case: Adaptive Multilevel Splitting (AMS) algorithm (see after).

Adaptivity and Mutations

Bias and consistency

Indexing using selection

AMS 000000000000

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- Consistency of Adaptive mutations: large mixing $M_s \rightarrow \eta_s$.
- Well-known rare event case: Adaptive Multilevel Splitting (AMS) algorithm (see after).
- AMS in the dynamical setting has a hidden non-adaptive Feynman-Kac-Del Moral structure (see below).

Adaptivity and Mutations

Bias and consistency

Indexing using selection

Remark on Adaptive Mutations

Additional Intrinsic Adaptivity on Mutations: The mutation kernel M_s is random and depends on the past replica empirical distribution. E.g.: if M_s is based on accept/reject, proposal is adaptively tuned to target an average acceptance rate $r_0 \in (0, 1)$.

Adaptivity and Mutations

Bias and consistency ●000 Indexing using selection

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The Feynman-Kac-Del Moral structure

Adaptivity and Mutations

The Feynman-Kac-Del Moral structure

• For non-adaptive = preset mutations, the algorithm can be derived from a Feynman-Kac (FK) formula:

$$\int \varphi(x) e^{-V_{s_{(i)}}(x)} \pi(dx) = \\ \mathbb{E} \left[\varphi(X_{s_{(i)}}) e^{-\sum_{i'=1}^{i} V_{s_{(i')}}(X_{s_{(i'-1)}}) - V_{s_{(i'-1)}}(X_{s_{(i'-1)}})} \right]$$

where $X_{s_{(i)}}$, $i \ge 0$ is a Markov chain with $X_0 \sim \eta_0$ and probability transition $M_{s_{(i)}}$.

Adaptivity and Mutations

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• The algoritm is then: simulating independently *N* chains with FK weights. Additional re-sampling/selection to prevent weight degeneracy.

Adaptivity and Mutations

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- The algoritm is then: simulating independently *N* chains with FK weights. Additional re-sampling/selection to prevent weight degeneracy.
- Nota Bene: in Del Moral, re-sampling/selection is put in a (very slightly restrictive) 'mean-field' form.

Context and Algo	Adaptivity and Mutations	Bias and consistency 0●00	Indexing using selection	AMS 00000000000

Remark

The Feynman-Kac formula before is known in physics as 'Jarzynski equality'. In that case:

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Context and Algo	Adaptivity and Mutations	Bias and consistency 0●00	Indexing using selection	AMS 00000000000

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• *s* is reaction coordinate or a thermodynamic parameter.

Context and Algo	Adaptivity and Mutations	Bias and consistency 0●00	Indexing using selection	AMS 00000000000

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Context and Algo	Adaptivity and Mutations	Bias and consistency 0●00	Indexing using selection	AMS 00000000000

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Context and Algo	Adaptivity and Mutations	Bias and consistency 0●00	Indexing using selection	AMS 00000000000

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Context and Algo	Adaptivity and Mutations	Bias and consistency 0●00	Indexing using selection	AMS 00000000000

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- Weight = $e^{-Work/(k_bT)}$!!
- Exists experimentally !!

Adaptivity and Mutations

Bias and consistency

Indexing using selection

AMS 000000000000

The Feynman-Kac-Del Moral structure

Proposition (Unbiasedness)

Un-normalized estimators are unbiased for algorithms following the Feynman-Kac-Del Moral structure.

Proof.

First remark that $\int \varphi e^{-V_{s_{(i)}}} d\pi = \mathbb{E}[\varphi(X_{s_{(i)}})e^{-V_{s_{(i)}}(X_{s_{(i-1)}})+V_{s_{(i-1)}}(X_{s_{(i-1)}})} \times \ldots \times e^{-V_{s_{(1)}}(X_{s_{(0)}})+V_{s_{(0)}}(X_{s_{(0)}})}] =: \mathbb{E}[Q^{0 \to i}(\varphi)(X_{0})] \text{ where } i \mapsto X_{(i)} \text{ is the}$ MCMC chain used in the mutation step. Then check that for $i \leq i_{0}$

$$i \mapsto z_{s^{(i)}}^{N} \int Q^{i \to i^{0}}(\varphi) \, d\eta_{s^{(i)}}^{N}$$
 is a martingale.

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Adaptivity and Mutations

Bias and consistency

Indexing using selection $\bullet 000$

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High dimension requires continuous

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Bias and consistency

Indexing using selection

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High dimension requires continuous

• High Dimension $d \gg 1$: weights that are \times by $e^{-V_{s_{(i+1)}}(X_{s_{(i)}})+V_{s_{(i)}}(X_{s_{(i)}})}$ at each iteration have exponential variance with d (typically).

Example

In \mathbb{R}^d , if coordinates of X are i.i.d. and V has a sum form over coordinates and is smooth w.r.t. s, by CLT, non-degeneracy of weights requires:

$$s^{(i+1)} - s^{(i)} \sim rac{1}{\sqrt{d}} \stackrel{d o +\infty}{\longrightarrow} 0.$$

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Bias and consistency

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Context and Algo Adaption

Adaptivity and Mutations

Bias and consistency

Indexing using selection

AMS 00000000000

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- Tempting to not mutate at each $s^{(i)}$.
- Idea: <u>switch to a continuum of scores</u>:

$$s \in \left\{s^{(0)}, \dots, s^{(I)}
ight\}$$
 becomes $s \in [0, 1]$.

Adaptivity and Mutations

Bias and consistency

Indexing using selection $0 \bullet 00$

Indexing the algorithm by selection events

'Same' algorithm, new representation:



Adaptivity and Mutations

Bias and consistency

Indexing using selection 0000

Indexing the algorithm by selection events

'Same' algorithm, new representation:

• Preset Mutations: Each replicas evolve independently according to a Markov process with generator L_s invariant with respect to target $\eta_s \propto e^{-V_s} \pi$.

Example

Piecewise constant Markov jump process

$$L_s(\varphi)(x) = \lambda_s(M_s(\varphi)(x) - \varphi(x)), \quad \eta_s M_s = \eta_s$$

can be simulated: i) mutations occur at random score (higher than s_0 with proba $e^{-\int_0^{s_0} \lambda_s ds}$), ii) mutations with M_s .

Context and Algo Adaptiv

Adaptivity and Mutations

Bias and consistency

Indexing using selection $0 \bullet 00$

AMS 000000000000

Indexing the algorithm by selection events

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• Other examples: discretization of a Stochastic Differential Equation, or Piecewise Deterministic Markov Process.

Adaptivity and Mutations Bias and consistency

Indexing using selection 0000

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Re-Indexing the algorithm by splitting events

Initialize replicas and set $S_{(0)} = 0$. Mutate all particles with L_s on $s \in [0, 1]$. Iterate on *j*:

Adaptivity and Mutations

Bias and consistency

Indexing using selection 0000

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Initialize replicas and set $S_{(0)} = 0$. Mutate all particles with L_s on $s \in [0, 1]$. Iterate on j:

 (j) Weights: compute the 'importance' weight for s ∈ [0, 1] of replicas so that it targets η_s for each s, e.g.: e<sup>-∫₀^s ∂_{s'}V_{s'}(X_{s'})ds'.
</sup> Context and Algo Adaptivity and Mutations

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- (*j*) Selection Compute the next random score

$$S_{(j)} := \inf \left\{ s \geqslant S_{(j-1)} | \mathsf{Criteria}_s^N == 1
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e.g.: Criteria_s = weight degeneracy (Effective Sample Size) at s.

Then perform selection/re-sampling according to weights.

Adaptivity and Mutations Context and Algo

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- (j) Preset Mutations: mutate with L_s on $s \in [S^{(j)}, 1]$ new (\Leftrightarrow all !) replicas.

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Adaptivity and Mutations

Bias and consistency

Indexing using selection 0000

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- (j) Preset Mutations: mutate with L_s on $s \in [S^{(j)}, 1]$ new (\Leftrightarrow all !) replicas.
- (Exit) Stop if $S^{(j)} = 1$ else $j \to j + 1$.

Adaptivity and Mutations

Bias and consistency

Indexing using selection 0000

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Re-Indexing the algorithm by splitting events

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Adaptivity and Mutations

Bias and consistency

Indexing using selection $000 \bullet$

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Re-Indexing the algorithm by splitting events

Remarks

• Preset mutations are simulated by ANTICIPATION (can be adjusted to decrease cost).

Adaptivity and Mutations

Bias and consistency

Indexing using selection $000 \bullet$

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- Unbiasedness/Feynman-Kac/Del Moral structure^a holds if no Triggered-Mutation .

Adaptivity and Mutations

Bias and consistency

Indexing using selection $000 \bullet$

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- Unbiasedness/Feynman-Kac/Del Moral structure^a holds if no Triggered-Mutation .
- AMS in 'static setting' is an example with ONLY Triggered Mutations-On-Child (see after).

Adaptivity and Mutations

Bias and consistency

Indexing using selection $000 \bullet$

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Re-Indexing the algorithm by splitting events

Remarks

- Preset mutations are simulated by ANTICIPATION (can be adjusted to decrease cost).
- Unbiasedness/Feynman-Kac/Del Moral structure^a holds if no Triggered-Mutation .
- AMS in 'static setting' is an example with ONLY Triggered Mutations-On-Child (see after).
- AMS in 'dynamic setting' is an example with PSEUDO-triggered Mutation-On-Child: they are in fact preset mutations given by the model itself !, (see after).

^aSee also Brehier Gazeau Goudenege Lelievre Rousset GAMS 2016

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS •0000000000
Static ² AM	S algorithm			

Let k < N given. Assume rare event setting with:

²F Cérou, P Del Moral, T Furon, A Guyader Sequential Monte Carlo for rare event estimation 2012

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS •00000000000
Static ² AM	1S algorithm			

Let k < N given. Assume rare event setting with:

• $\pi :=$ anything simulable.

²F Cérou, P Del Moral, T Furon, A Guyader Sequential Monte Carlo for rare event estimation 2012
Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS •0000000000
Static ² AM	S algorithm			

- $\pi :=$ anything simulable.
- $e^{-V_s} = \mathbf{1}_{score>s}$.

²F Cérou, P Del Moral, T Furon, A Guyader *Sequential Monte Carlo for* rare event estimation 2012

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- $\pi :=$ anything simulable.
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- Selection = killing + neutral bearing. Triggered by k replicas with lowest score which are killed and then neutrally borne.
- Mutation-If-Selection with Mutation-On-Child. M_s is a MCMC kernel reversible w.r.t. π with rejection if proposal has score $\leq s$.

²F Cérou, P Del Moral, T Furon, A Guyader *Sequential Monte Carlo for* rare event estimation 2012

Adaptivity and Mutations

Bias and consistency

Indexing using selection 0000

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Dynamical³ AMS algorithm

³F Cérou, A Guyader, Adaptive multilevel splitting for rare event analysis 📱 🗠 🤉

Adaptivity and Mutations

Bias and consistency

Indexing using selection

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Dynamical³ AMS algorithm

• $\pi = Law$ of a Markov chain / process.

³F Cérou, A Guyader, Adaptive multilevel splitting for rare event analysis 🚊 🗠 🔍

Adaptivity and Mutations

Bias and consistency

Indexing using selection 0000

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Dynamical³ AMS algorithm

π = Law of a Markov chain / process.
e^{-V_s} = 1_{score>s}, score = max(ζ(path)).

³F Cérou, A Guyader, Adaptive multilevel splitting for rare event analysis 📱 🗠 🤉

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³F Cérou, A Guyader, Adaptive multilevel splitting for rare event analysis 🛓 🕤 🖉

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Indexing using selection

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- The algorithm is just a Fleming-Viot process indexed by s.

³F Cérou, A Guyader, Adaptive multilevel splitting for rare event analysis 🚊 🗠 🔍

Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS ००●०००००००

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Adaptive Multilevel Splitting

- Black line: $\{\xi = constant\}.$



Adaptivity and Mutations

Bias and consistency

Indexing using selection

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Adaptivity and Mutations

Bias and consistency

Indexing using selection



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Context and Algo	Adaptivity and Mutations 0 00000	Bias and consistency	Indexing using selection	AMS 0000000000
Fluctuatio	n Analysis			

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Context and Algo	Adaptivity and Mutations	Bias and consistency	Indexing using selection	AMS 0000000000
Fluctuation	Analysis			

• Thanks to the FK-Del Moral structure, we can carry out martingale analysis an obtain Central Limit Theorem (Warning: practical issues with CLT).

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Fluctuation	Analysis			

- Thanks to the FK-Del Moral structure, we can carry out martingale analysis an obtain Central Limit Theorem (Warning: practical issues with CLT).
- For dynamic AMS, e.g. estimator of rare event is asymptotically normal when $N \rightarrow +\infty$ with asymptotic variance:

$$\operatorname{Var}(\hat{\rho}) = -p^2 \ln \rho + 2 \int_{s \ge 0} \operatorname{Var}_{\eta_s}(\mathbb{P}_{\cdot}(\operatorname{rare\ event})) d(-p_s^2)$$

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Variance is Minimal = -p² ln p for diffusions iff P_x(rare event) only depends on x only through ξ(x).

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- Variance is Minimal = -p² ln p for diffusions iff P_x(rare event) only depends on x only through ξ(x).
- Ref: Cérou, (Delyon), Guyader , Rousset: series of paper.