

Branching Distributional Equations and their Applications

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Motivation

- ▶ Two problems leading to branching distributional equations:
 - ▶ Ranking algorithms on complex networks.
 - ▶ Parallel queueing systems with synchronization.

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- ▶ The object of interest is the **attracting endogenous** solution to a certain max-plus branching stochastic fixed-point equation.

Motivation

- ▶ Two problems leading to branching distributional equations:
 - ▶ Ranking algorithms on complex networks.
 - ▶ Parallel queueing systems with synchronization.
- ▶ The object of interest is the **attracting endogenous** solution to a certain max-plus branching stochastic fixed-point equation.
- ▶ Naïve Monte Carlo techniques are not a good idea...
- ▶ **Goal:** provide an efficient simulation algorithm to study them.

Example: PageRank

- ▶ Consider a directed graph on n vertices.
- ▶ Generalized PageRank assigns to each vertex i a “rank” r_i satisfying:

$$r_i = q_i + \sum_{j \rightarrow i} \frac{\zeta_j}{d_j} \cdot r_j,$$

where q_i is known as the personalization or teleportation value of vertex i , d_j is the out-degree of vertex j , $|\zeta_j| \leq c < 1$, and the sum is taken over all inbound neighbors of vertex i .

- ▶ Multiplying both sides by n , we obtain the “scale-free” ranks $R_i = nr_i$.
- ▶ The goal is to compute the limiting distribution of the scale-free rank of a randomly chosen vertex.

Example: PageRank

- ▶ Consider random digraphs constructed according to either the directed configuration model or the inhomogeneous random digraph with rank-1 kernel.
- ▶ The distribution of the scale-free rank of a randomly chosen vertex, R_ξ , converges, as $n \rightarrow \infty$, to

$$R^* = Q_0 + \sum_{j=1}^{N_0} X_j,$$

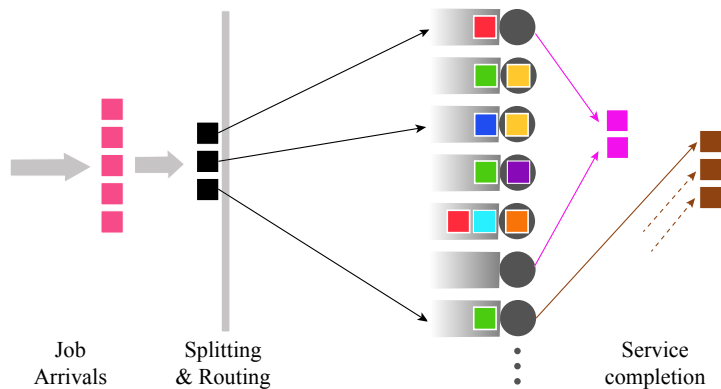
where the $\{X_j\}$ are i.i.d. copies of the attracting endogenous solution to the linear SFPE:

$$X \stackrel{\mathcal{D}}{=} QC + \sum_{j=1}^N CX_j$$

Example: Queueing networks

- ▶ Consider a network consisting of n computers working in parallel, e.g. a server farm.
- ▶ Jobs arrive to the network according to a Poisson process.
- ▶ Upon arrival, they are split into several pieces of similar size, with larger jobs having more pieces.
- ▶ The pieces are then routed to a random subset of servers to be processed separately.
- ▶ **Network rules:**
 - ▶ Jobs need to be served on a FCFS basis.
 - ▶ All pieces of a job must begin service simultaneously (**Synchronization**).
- ▶ The synchronization causes idleness and blocking.

Example: Queueing networks



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- ▶ We study the distribution of the stationary waiting time of a job in the n server network, denoted $W^{(n)}$.
- ▶ As the number of servers $n \rightarrow \infty$, $W^{(n)}$ converges to the high-order Lindley equation:

$$W \stackrel{\mathcal{D}}{=} \max \left\{ 0, \max_{1 \leq i \leq N} (\chi_i - \tau_i + W_i) \right\},$$

where τ_i is exponentially distributed, N denotes the number of fragments of a job, and the $\{\chi_i\}$ are the service requirements of each of the N fragments.

Solutions to branching recursions

- ▶ The attracting endogenous solutions to

$$R \stackrel{\mathcal{D}}{=} Q + \sum_{i=1}^N C_i R_i \quad \text{and} \quad R \stackrel{\mathcal{D}}{=} Q \vee \left(\bigvee_{i=1}^N C_i R_i \right)$$

are especially simple, and are given by

$$R = \sum_{\mathbf{i} \in \mathcal{T}} \Pi_{\mathbf{i}} Q_{\mathbf{i}} \quad \text{and} \quad R = \bigvee_{\mathbf{i} \in \mathcal{T}} \Pi_{\mathbf{i}} Q_{\mathbf{i}},$$

respectively, where the sum (maximum) is taken over all nodes $\mathbf{i} = (i_1, \dots, i_n)$ on the tree \mathcal{T} .

Finite tree approximation

- ▶ Often, the map Φ defines a contraction under a suitable metric.
- ▶ More precisely, let

$$R_{\mathbf{i}}^{(j)} = \Phi \left(Q_{\mathbf{i}}, N_{\mathbf{i}}, \{C_{(\mathbf{i},r)}\}, \{R_{(\mathbf{i},r)}^{(j-1)}\} \right), \quad 1 \leq j \leq k,$$

for \mathbf{i} in generation $k - j$, and $R_{\mathbf{i}}^{(0)} = Q_{\mathbf{i}}$ for \mathbf{i} in generation k .

- ▶ Define μ_k to be the probability measure of $R^{(k)}$ and μ that of the attracting endogenous solution R .
- ▶ In many cases we can show that there exist a metric d and a constant $0 < c < 1$ such that

$$d(\mu_k, \mu) \leq c^k d(\mu_0, \mu).$$

- ▶ Hence, it suffices to compute the distribution of $R^{(k)} = R_{\emptyset}^{(k)}$ for large k .

A simulation approach

- ▶ In principle, we could simulate a tree for k generations (k large) and use

$$R \approx R^{(k)}.$$

- ▶ For the linear and maximum recursion this would imply computing

$$R^{(k)} = \sum_{\mathbf{i} \in \mathcal{T}, |\mathbf{i}| \leq k} \Pi_{\mathbf{i}} Q_{\mathbf{i}} \quad \text{and} \quad R^{(k)} = \bigvee_{\mathbf{i} \in \mathcal{T}, |\mathbf{i}| \leq k} \Pi_{\mathbf{i}} Q_{\mathbf{i}},$$

where $|(i_1, \dots, i_n)| = n$.

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where $|(i_1, \dots, i_n)| = n$.

- ▶ **Problem:** This approach requires an average of $(E[N])^k$ i.i.d. copies of $(Q, N, \{C_i\})!$

The Population Dynamics algorithm

1. **Initialize:** Set $j = 0$. Simulate a sequence $\{Q_i\}_{i=1}^m$ of i.i.d. copies of Q and let $\hat{R}_i^{(0,m)} = Q_i$ for $i = 1, \dots, m$. Output $\mathcal{P}^{(0,m)} = \left(\hat{R}_1^{(0,m)}, \hat{R}_2^{(0,m)}, \dots, \hat{R}_m^{(0,m)} \right)$ and update $j = 1$.

2. While $j \leq k$:

2.1 Simulate a sequence $\{(Q_i, N_i, C_{(i,1)}, C_{(i,2)}, \dots)\}_{i=1}^m$ of i.i.d. copies of the generic branching vector, independent of everything else.

2.2 Let

$$\hat{R}_i^{(j,m)} = \Phi \left(Q_i, N_i, \{C_{(i,r)}\}, \{\hat{R}_{(i,r)}^{(j-1,m)}\} \right), \quad i = 1, \dots, m, \quad (1)$$

where the $\hat{R}_{(i,r)}^{(j-1,m)}$ are sampled uniformly with replacement from the pool $\mathcal{P}^{(j-1,m)}$.

2.3 Output $\mathcal{P}^{(j,m)} = \left(\hat{R}_1^{(j,m)}, \hat{R}_2^{(j,m)}, \dots, \hat{R}_m^{(j,m)} \right)$ and update $j = j + 1$.

Estimators

- ▶ For large enough m , we expect that

$$\hat{F}_{k,m}(x) \triangleq \frac{1}{m} \sum_{i=1}^m 1(\hat{R}_i^{(k,m)} \leq x) \approx P(R^{(k)} \leq x) \triangleq F_k(x).$$

- ▶ Moreover, for functions h “nice enough”, we would want to estimate $E[h(R^{(k)})]$ using

$$\frac{1}{m} \sum_{i=1}^m h(\hat{R}_i^{(k,m)}).$$

- ▶ **Problem:** Although the $\{\hat{R}_i^{(k,m)}\}$ are identically distributed, they are not independent and do not in general have distribution F_k .

Assumptions

- ▶ We have two different types of assumptions:
- ▶ **Assumption p :** Suppose that for some $p \geq 1$ there exists a constant $0 < H_p < \infty$ such that

$$E [|\Phi(Q, N, \{C_i\}, \{X_i\}) - \Phi(Q, N, \{C_i\}, \{Y_i\})|^p] \leq H_p E [|X_1 - Y_1|^p],$$

for $\{(X_i, Y_i)\}$ i.i.d., independent of $(Q, N, \{C_i\})$, and such that $\{X_i\} \sim F$ and $\{Y_i\} \sim G$.

- ▶ **Assumption φ :** There exists a function $\varphi : \mathbb{R} \rightarrow \mathbb{R}_+$ such that

$$|\Phi(q, n, \{c_r\}, \{x_r\}) - \Phi(q, n, \{c_r\}, \{y_r\})| \leq \sum_{r=1}^n \varphi(c_r) |x_r - y_r|$$

for any vectors $(q, n, \{c_r\}, \{x_r\})$ and $(q, n, \{c_r\}, \{y_r\})$ for which Φ is well-defined.

The Wasserstein metric

- ▶ **Definition:** Let F and G be distribution functions on \mathbb{R} having finite p th moment, $p \geq 1$. The Wasserstein distance of order p between F and G is

$$\begin{aligned}d_p(F, G) &= \inf_{X \sim F, Y \sim G} (E[|X - Y|^p])^{1/p} \\ &= \left(\int_0^1 |F^{-1}(u) - G^{-1}(u)|^p du \right)^{1/p},\end{aligned}$$

where $f^{-1}(t) = \inf\{x \in \mathbb{R} : f(x) \geq t\}$.

- ▶ Convergence in the Wasserstein metric of order p is equivalent to weak convergence plus convergence of the absolute moments of order up to p .
- ▶ Recall that

$$F_k(x) = P(R^{(k)} \leq x) \quad \text{and} \quad \hat{F}_{k,m}(x) = \frac{1}{m} \sum_{i=1}^m 1(\hat{R}_i^{(k,m)} \leq x).$$

Convergence in mean

- **Theorem A:** Fix $1 \leq p < \infty$ and suppose that Φ satisfies *Assumption p* for p . Assume further that for any fixed $k \in \mathbb{N}$, $\max_{0 \leq j \leq k} E[(R^{(j)})^p] < \infty$. Then,

$$E \left[d_p(\hat{F}_{k,m}, F_k)^p \right] \leq \left(\sum_{r=0}^k (H_p^{1/p})^r \right)^{p-1} \sum_{r=0}^k (H_p^{1/p})^{k-r} E \left[d_p(F_{r,m}, F_r)^p \right],$$

where $F_{r,m}$ is the empirical distribution function of an i.i.d. sample of size m from F_r . Moreover, if $\max_{0 \leq j \leq k} E[|R^{(j)}|^q] < \infty$ for $q > p \geq 1$, $q \neq 2p$, then

$$E \left[d_p(\hat{F}_{k,m}, F_k)^p \right] \leq K m^{-\alpha} \left(\sum_{r=0}^k (H_p^{1/p})^r \right)^{p-1} \sum_{r=0}^k (H_p^{1/p})^{k-r} (E[|R^{(j)}|^q])^{p/q},$$

where $K = K(p, q)$ and $\alpha = \min\{(q - p)/q, 1/2\}$.

Some remarks

- ▶ The theorem does not require $H_p < 1$.
- ▶ When $H_p < 1$, the bound provided by the theorem becomes independent of the depth of the recursion k .
- ▶ In some cases, failure to find $H_p < 1$ for some $p \geq 1$ implies the instability of the recursion.
- ▶ The rate at which $E[d_p(\hat{F}_{k,m}, F_k)^p]$ converges to zero is that of $\max_{0 \leq j \leq k} E[d_p(F_{j,m}, F_j)^p]$ (up to a constant), which corresponds to the “perfect” i.i.d. pool.

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There is no loss in the rate of convergence due to the bootstrapping!

Almost sure convergence

- **Theorem B:** Fix $1 \leq p < \infty$ and suppose that Φ satisfies *Assumption p* for both p and $2p$. Assume further that for any fixed $k \in \mathbb{N}$, $\max_{0 \leq j \leq k} E[(R^{(j)})^{2p}(\log |R^{(j)}|)^+] < \infty$. Then,

$$\lim_{m \rightarrow \infty} d_p(\hat{F}_{k,m}, F_k) = 0 \quad \text{a.s.}$$

- **Theorem C:** Fix $1 \leq p < \infty$ and suppose that Φ satisfies *Assumption φ* . Assume further that $[|R^{(0)}|^{p+\delta} + Z^{p+\delta}] < \infty$ for some $\delta > 0$, where $Z = \sum_{i=1}^N \varphi(C_i)$. Then, for any fixed $k \in \mathbb{N}$,

$$\lim_{m \rightarrow \infty} d_p(\hat{F}_{k,m}, F_k) = 0 \quad \text{a.s.}$$

Consistency of estimators

- ▶ **Definition:** We say that Θ_n is a *weakly consistent* estimator for θ if, as $n \rightarrow \infty$, $\Theta_n \xrightarrow{P} \theta$, and we say that it is *strongly consistent* if $\Theta_n \rightarrow \theta$ a.s.
- ▶ **Proposition:** Fix $1 \leq p < \infty$ and suppose that $h : \mathbb{R} \rightarrow \mathbb{R}$ satisfies $|h(x)| \leq C(1 + |x|^p)$ for all $x \in \mathbb{R}$ and some constant $C > 0$. Then,
 - i.) If $E[d_p(\hat{F}_{k,m}, F_k)^p] \rightarrow 0$ as $m \rightarrow \infty$, then

$$\frac{1}{m} \sum_{i=1}^m h(\hat{R}_i^{(k,m)}) = \int_{-\infty}^{\infty} h(x) d\hat{F}_{k,m}(x) \quad (2)$$

is a weakly consistent estimator for $E[h(R^{(k)})]$ for each fixed $k \in \mathbb{N}$.

- ii.) If $d_p(\hat{F}_{k,m}, F_k) \rightarrow 0$ a.s. as $m \rightarrow \infty$, then (2) is a strongly consistent estimator for $E[h(R^{(k)})]$ for each fixed $k \in \mathbb{N}$.

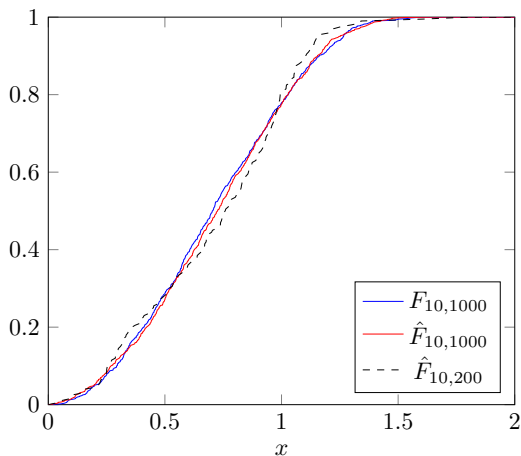
Open problem

- ▶ Derive a CLT for the algorithm.
- ▶ *Hint:* Suppose that Assumption p holds for some $1 \leq p < \infty$, and that $\max_{0 \leq r \leq k} E[|R^{(r)}|^q] < \infty$ for some $q > 2p$.
- ▶ Under these conditions **Theorem B** holds and it should be easy to verify that

$$\frac{1}{m} \sum_{i=1}^m h(\hat{R}_i^{(k,m)}) - E \left[h(\hat{R}_1^{(k,m)}) \mid \mathcal{P}^{(k-1,m)} \right],$$

properly scaled, satisfies a CLT.

A numerical example (Linear SFPE)



$F_{10,1000}$ is the empirical CDF of $R^{(10)}$ with a sample of 1000 (computation time 883 sec.); $\hat{F}_{10,1000}$ required only 2.1 sec.