

# Subexponential LD & Condensing ZRP

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joint work with Inés Armendáriz and Stefan Grosskinsky

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# A toy example

Suppose  $\{X_n\}_{n \in \mathbb{N}}$  is a sequence of i.i.d. r.v.'s with

$$X_i \sim \mathcal{N}(m, \sigma^2).$$

Then,  $\bar{X}_n := \frac{X_1 + \cdots + X_n}{n} \sim \mathcal{N}(m, \frac{\sigma^2}{n})$

and for any  $x > m$

$$\mathbb{P}[\bar{X}_n \geq x] = \int_{\frac{\sqrt{n}(x-m)}{\sigma}}^{\infty} e^{-\frac{u^2}{2}} \frac{du}{\sqrt{2\pi}} \sim \frac{\sigma}{\sqrt{2\pi n}(x-m)} e^{-\frac{n(x-m)^2}{2\sigma^2}}.$$

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That is,

$$\frac{1}{n} \log \mathbb{P}[\bar{X}_n \geq x] \longrightarrow -\frac{(x-m)^2}{2\sigma^2}.$$

# A toy example II

Suppose  $\{X_n\}_{n \in \mathbb{N}}$  is a sequence of i.i.d. r.v.'s with

$$X_i \sim \mathcal{N}(m, \sigma^2).$$

Then,

$$X_i - \bar{X}_n \sim \mathcal{N}\left(0, \frac{(n-1)\sigma^2}{n}\right)$$

and

$$X_i - \bar{X}_n, \bar{X}_n \text{ are independent.}$$

Hence,

$$\mathcal{L}[X_i \mid \bar{X}_n = x] = \mathcal{N}\left(x, \frac{(n-1)\sigma^2}{n}\right) \longrightarrow \mathcal{N}(x, \sigma^2).$$

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With a bit more work we find

$$\text{Cov}(X_i - \bar{X}_n, X_j - \bar{X}_n) = \sigma^2 \delta_{ij} - \frac{\sigma^2}{n}$$

and, if we set  $\mathbf{e}^\top = (1, 1, \dots, 1)$ , then for any  $k \in \mathbb{N}$

$$\mathcal{L}[(X_1, \dots, X_k) \mid \bar{X}_n = x] = \mathcal{N}(x\mathbf{e}, \sigma^2 \mathbb{I}_k - \frac{\sigma^2}{n} \mathbf{e} \mathbf{e}^\top) \longrightarrow \mu^k$$

where the law  $\mu$  is  $\mathcal{N}(x, \sigma^2)$ .

# Cramér's theorem

Suppose  $\{X_n\}_{n \in \mathbb{N}}$  are i.i.d. r.v.'s with common law  $\mu$  and

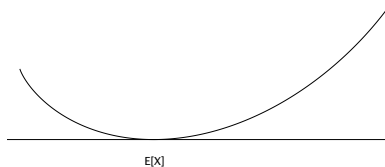
$$M(\lambda) = \mathbb{E}[e^{\lambda X_i}] < \infty, \text{ for } |\lambda| \leq \lambda_0.$$

If  $x > \mathbb{E}[X_i]$ , then

$$\frac{1}{n} \log \mathbb{P}[\bar{X}_n \geq x] \longrightarrow -I(x),$$

where

$$I(x) = \sup_{\lambda} (\lambda x - \log M(\lambda))$$



[Cramér 1938].

# Gibbs conditioning

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$\mu_*$  solves the following variational problem

$$\inf_{\nu \in \mathcal{I}} H(\nu \mid \mu), \quad \text{where} \quad \mathcal{I} = \left\{ \nu \in M_1(\mathbb{R}) : \int u \nu(du) \geq x \right\}$$

and

$$H(\nu \mid \mu) = \begin{cases} \int f \log f \, d\mu & \text{if } \nu \ll \mu \text{ with } f = \frac{d\nu}{d\mu}, \\ +\infty & \text{otherwise} \end{cases}$$



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We may take  $k = k(n) \rightarrow \infty$  slowly, but the result is no longer true if  $k = O(n)$ .

[Dembo-Zeitouni 1996].

# Subexponential Distributions

The picture is completely different when  $\mu$  has no exponential moments, i.e.  $M(\lambda) = \mathbb{E}[e^{\lambda X_i}] = \infty$  for all  $\lambda > 0$ . A distribution  $\mu$  supported on the positive half-line is called subexponential if

$$\lim_{x \rightarrow \infty} \frac{\mathbb{P}[X + Y > x]}{\mathbb{P}[X > x]} = 2,$$

where  $X, Y$  are independent  $\mu$ -distributed r.v.'s.

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Heavy-tailed distributions typically used in applications are all in this class.

- Regularly varying tails:  $\mathbb{P}[X_1 > x] \sim x^{-\gamma} L(x)$  with  $\gamma > 0$ ,
- Lognormal type tails:  $\mathbb{P}[X_1 > x] \sim x^{-\beta} e^{-\gamma(\log x)^\lambda}$ ,  $\lambda > 1$ .
- Weibull type tails:  $\mathbb{P}[X_1 > x] \sim x^{-\beta} e^{-\gamma x^\lambda}$ ,  $\lambda < 1$ .

# Large Deviation Probabilities

When  $X_1, X_2, \dots, X_n$  are i.i.d. and subexponential the large deviations probabilities of their sum are typically given by

$$\mathbb{P}[X_1 + \dots + X_n > x] \sim n\mathbb{P}[X_1 > x] \quad [\text{Heyde 1968, Nagaev 1969, \dots, Denisov et al 2009}]$$

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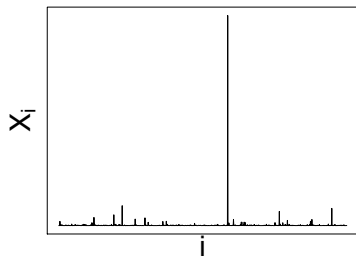
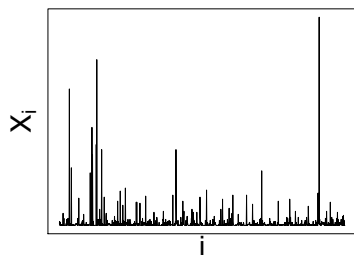
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Note that since it is always true that

$$\mathbb{P}[\max\{X_1, \dots, X_n\} > x] \sim n\mathbb{P}[X_1 > x],$$

subexponentiality implies that a large deviation of the sum is typically realised by a single big jump.



# Gibbs conditioning for Subexponential r.v.'s

## Theorem (I. Armendáriz, ML)

Let  $X_1, X_2, \dots$  be i.i.d. r.v.'s with subexponential distribution  $\mu$ . Define  $\mu_x = \mathcal{L}[X_i \mid X_i > x]$ , and

$$\mu_{n,x} = \mathcal{L}[(X_1, \dots, X_n) \mid X_1 + \dots + X_n > x].$$

Then,

$$\lim_{x \rightarrow \infty} \sup_{n \leq A(x)} \left\| \mu_{n,x} - \frac{1}{n} \sum_{j=1}^n \sigma^j(\mu^{n-1} \times \mu_x) \right\|_{\text{t.v.}} = 0.$$

The maximum entirely absorbs the correlations introduced by conditioning- the bulk becomes asymptotically independent.

# Idea of the proof

**Proof:** (for nonnegative r.v.'s) Note that if  $\mu$  is a probability measure,  $\mu[A] > 0$ , and  $\mu_A[\cdot] = \mu[\cdot | A]$ , then  $\mu_A$  is the solution to the minimization problem

$$\min_{\nu[A]=1} H(\nu|\mu).$$

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Candidate distribution:  $\mu_{n,x}^* = \frac{1}{n} \sum_{j=1}^n \sigma^j(\mu^{n-1} \times \mu_x)$ . By Csiszár's parallelogram identity and Pinsker's inequality

$$\begin{aligned} \|\mu_{n,x} - \mu_{n,x}^*\|_{\text{t.v.}}^2 &\leq H(\mu_{n,x}^* | \mu^n) - H(\mu_{n,x} | \mu^n) \\ &\leq \log \left( \frac{\mathbb{P}[S_n > x]}{n\mathbb{P}[X_1 > x]} \right) + n\mathbb{P}[X_1 > x]. \end{aligned}$$

# Gibbs conditioning for Subexponential r.v.'s (local case)

## Theorem

Let  $X_1, X_2 \dots$  be i.i.d. lattice r.v.'s with subexponential distribution  $\mu$ . For admissible values of  $x$ , define

$$\mu_{n,x} = \mathcal{L}[(X_1, \dots, X_n) \mid S_n = x].$$

If  $\nu_{n,x}^j$  is a distribution on  $\mathbb{R}^n$  with marginal on the co-ordinates other than  $j$  equal to  $\mu^{n-1}$  and conditional distribution of the  $j$ -th co-ordinate given the others  $\delta_{x - \sum_{i \neq j} x_i}$  then

$$\lim_{x \rightarrow \infty} \sup_{n \leq A(x)} \left\| \mu_{n,x} - \frac{1}{n} \sum_{j=1}^n \nu_{n,x}^j \right\|_{\text{t.v.}} = 0.$$

# The condensing zero-range process

**State space:**  $\mathbb{X} = \{0, 1, \dots\}^\Lambda$      $\eta = (\eta_x)_{x \in \Lambda}$ .

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**State space:**  $\mathbb{X} = \{0, 1, \dots\}^\Lambda$      $\eta = (\eta_x)_{x \in \Lambda}$ .

**Dynamics:** If there are  $k$  particles at a site  $x$ , one of them leaves after an exponential time with rate  $g(k)$ , where

$$g : \{0, 1, 2, \dots\} \rightarrow [0, \infty)$$

and goes to  $y \in \mathbb{T}$  with probability  $p(x, y)$ .

[Spitzer, 1970]

**Jump rates:**  $g \downarrow$  : effective attraction.

A standard model for condensation

[Evans, 2000]

$$\begin{cases} g(k) = 1 + \frac{b}{k^\lambda} & k \in \mathbb{N} \\ g(0) = 0 \end{cases} \quad \text{for } \lambda \in (0, 1], \text{ (} b > 2 \text{ if } \lambda = 1.)$$

**Jump probabilities:**  $p(x, y) \in [0, 1]$

$$\sum_y p(x, y) = \sum_x p(x, y) = 1, \text{ walk irreducible.}$$

# Invariant product measures

**Conservation of the number of particles**  $\sum_{x \in \Lambda} \eta_x(t) = \text{const.}$

leads to a family of invariant product measures

**Grand-canonical measures** (fugacity  $\phi$ ) Product measures over  $\Lambda$  with marginals for  $\eta_x$

$$\nu_\phi[k] = \frac{1}{z(\phi)} \frac{\phi^k}{g!(k)} \quad \text{where} \quad g!(k) = \prod_{n=1}^k g(n)$$

defined when the partition function  $z(\phi) = \sum_k \frac{\phi^k}{g!(k)}$  converges.

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$$\frac{1}{g!(k)} \sim \begin{cases} k^{-b} & , \lambda = 1, b > 2 \\ \exp(-\frac{b}{1-\lambda} k^{1-\lambda}) & , \lambda \in (0, 1) \end{cases}$$

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**Density**  $R(\phi) = \langle \eta_x \rangle_{\nu_\phi} = \phi \partial_\phi \log z(\phi) \quad \uparrow \quad \text{in } \phi$

**Critical density**  $\rho_c = \lim_{\phi \nearrow \phi_c} R(\phi) < +\infty.$

# Canonical ensembles

Now consider the ZRP with  $N$  particles on  $|\Lambda_L| = L$  sites.

The process is irreducible over

$$X_{L,N} = \left\{ \boldsymbol{\eta} \in X_L : S_L(\boldsymbol{\eta}) := \sum_{x \in \Lambda_L} \eta_x = N \right\}.$$

**Canonical measures** Invariant measures  $\mu_{L,N}$  supported on  $X_{L,N}$

$$\mu_{L,N}[\boldsymbol{\eta}] = \frac{\prod_{x \in \Lambda_L} \frac{1}{g!(\eta_x)}}{\sum_{\boldsymbol{\eta}: S_L(\boldsymbol{\eta})=N} \prod_{x \in \Lambda_L} \frac{1}{g!(\eta_x)}}$$

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with  $\phi \leq \phi_c = 1$  and

$$\nu_{\phi}[k] \sim \frac{\phi^k}{g!(k)} = \phi^k \times \begin{cases} k^{-b} & \text{if } \lambda = 1, b > 2 \\ e^{-\frac{b}{1-\lambda}} k^{1-\lambda} & \text{if } \lambda \in (0, 1) \end{cases} .$$

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$$\mu_{L,N}[\cdot] = \nu_{\phi}^L \left[ \cdot \mid S_L(\boldsymbol{\eta}) = N \right]$$

**Question:** How large is  $M_L = \max_{x \in \Lambda_L} \eta_x$  under  $\mu_{L,N}$ ?

# Subcritical Densities

If  $N/L \rightarrow \rho < \rho_c$  there exists a fugacity  $\phi(\rho) < \phi_c (= 1)$  such that  $\langle \eta_x \rangle_{\nu_{\phi(\rho)}} = \rho$ .

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Think of  $\mu_{L,N}[\cdot] = \nu_{\phi(\rho)}^L \left[ \cdot \mid S_L(\boldsymbol{\eta}) = N \right]$ .

The event we are conditioning upon is not so unlikely, and locally  $\mu_{L,N}$  behaves as a product of  $\nu_{\phi(\rho)}$  in the limit (equivalence of ensembles.)

With a bit more work one sees that the typical size of  $M_L$  under  $\mu_{L,N}$  is the same as under  $\nu_{\phi(\rho)}^L$ .

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$$\mu_{L,N} \left[ \left| \frac{M_L}{\log L} - c(\rho) \right| > \epsilon \right] \rightarrow 0, \quad 0 < c(\rho) < +\infty.$$

# Supercritical Densities

If  $N/L \rightarrow \rho > \rho_c$  there exists no fugacity that corresponds to this density.

Now think of  $\mu_{L,N}[\cdot] = \nu_{\phi_c}^L \left[ \cdot \mid S_L(\boldsymbol{\eta}) = N \right]$ .

The event we are conditioning upon is a rare event and we need to understand how this large deviation of the sum is typically realised.

Recall that  $\nu_{\phi_c}$  is subexponential:

$$\nu_{\phi_c}[k] \sim \begin{cases} k^{-b} & \lambda = 1, b > 2 \\ e^{-\frac{b}{1-\lambda} k^{1-\lambda}} & \lambda \in (0, 1) \end{cases}$$

Understanding the invariant measures  $\mu_{L,N}$  reduces to Gibbs conditioning for subexponential r.v's

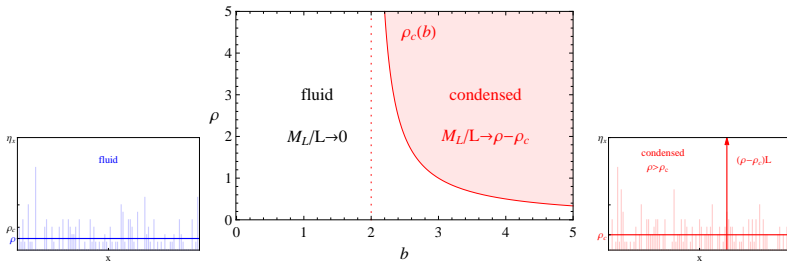
$$\mu_{L,N}[\cdot] = \nu_{\phi_c} \left[ \cdot \mid \sum_{x \in \mathbb{T}} \eta_x = N \right]$$

# Canonical measures and condensation

Equivalence of ensembles [Grosskinsky, Schütz, Spohn '03]

In the thermodynamic limit  $L, N \rightarrow \infty$ ,  $N/L \rightarrow \rho$

$$\mu_{L,N} \xrightarrow{w} \nu_\phi \quad \text{where} \quad \begin{cases} R(\phi) = \rho, & \rho \leq \rho_c \\ \phi = \phi_c, & \rho \geq \rho_c \end{cases} .$$



# Corollaries to the strong invariance principle

Let  $M_L = \max_{x \in \mathbb{T}} \eta_x$ .

- ① Since  $M_L +$  mass in the bulk =  $N$ ,  $\mu_{L,N}$ -a.s we get a conditional stable LT for the maximum from the stable LT for i.i.d. variables. If  $\nu_{\rho_c}$  has finite variance ( $b > 3$ )

$$\mu_{L,N} \left[ \frac{M_L - (N - \rho_c L)}{\sigma \sqrt{L}} \leq x \right] \rightarrow \int_{-\infty}^x e^{-y^2/2} \frac{dy}{\sqrt{2\pi}}.$$

(confirming the conjecture obtained from numerical simulations by Godrèche & Luck.)

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- ② Conditional  $k$ -order statistics converge to unconditional  $(k-1)$ -order statistics. E.g. if  $M_L^{(2)}$  is the second largest of the  $\{\eta_x\}_{x \in \mathbb{T}}$  then

$$\mu_{L,N} [M_L^{(2)} \leq x (\Gamma(b)L)^{\frac{1}{b-1}}] \rightarrow e^{-x^{1-b}}.$$

# Around criticality

[Armendáriz, Grosskinsky, L, 2013]

**Subcritical density:** The maximum  $M_L(\boldsymbol{\eta}) = \max_{x \in \Lambda_L} \eta_x$  is  $O(\log L)$  with Gumbel fluctuations.

$$\mu_{L,N} \left[ \frac{M_L(\boldsymbol{\eta}) - \alpha(\rho) \log L}{\beta_L} \leq x \right] \rightarrow e^{-e^{-x}}.$$

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**Supercritical density:** The maximum  $M_L(\boldsymbol{\eta})$  is  $O(L)$  with gaussian fluctuations.

$$\mu_{L,N} \left[ \frac{M_L(\boldsymbol{\eta}) - (N - \rho_c L)}{\sigma \sqrt{L}} \leq x \right] \longrightarrow \Phi(x).$$

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[Armendáriz, Grosskinsky, L, 2013]

**Subcritical density:** The maximum  $M_L(\boldsymbol{\eta}) = \max_{x \in \Lambda_L} \eta_x$  is  $O(\log L)$  with Gumbel fluctuations.

$$\mu_{L,N} \left[ \frac{M_L(\boldsymbol{\eta}) - \alpha(\rho) \log L}{\beta_L} \leq x \right] \rightarrow e^{-e^{-x}}.$$

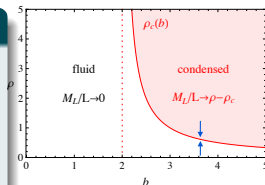
**Supercritical density:** The maximum  $M_L(\boldsymbol{\eta})$  is  $O(L)$  with gaussian fluctuations.

$$\mu_{L,N} \left[ \frac{M_L(\boldsymbol{\eta}) - (N - \rho_c L)}{\sigma \sqrt{L}} \leq x \right] \rightarrow \Phi(x).$$

## Question

How does  $M_L(\boldsymbol{\eta})$  behave as we go through the critical density? In particular, when does the condensate emerge?

Ferrari/ Evans, Majumdar 2008





# Small deviations

Let's focus on the case  $\lambda = 1$  and  $b > 3$  so that  $\sigma^2 < +\infty$ .

$$\nu_{\phi_c}[\eta_x = k] \sim k^{-b}.$$

There is a region around  $\rho_c L$  where the distribution of the maximum under  $\mu_{L,N}$  asymptotically behaves as the maximum of  $L$  independent samples drawn from  $\nu_{\phi_c}$ .

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## Proposition

If  $-L^{\frac{b-2}{b-1}} \ll N - \rho_c L \ll \sqrt{L \log L}$ , then

$$\mu_{L,N} \left[ M_L \leq x L^{\frac{1}{b-1}} \right] \sim \nu_{\phi_c}^L \left[ M_L \leq x L^{\frac{1}{b-1}} \right] \rightarrow e^{-u x^{1-b}} \quad \forall x > 0.$$

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**Critical behaviour:** Typical size of the maximum is  $L^{\frac{1}{b-1}}$  with Frechét fluctuations.

# Large Deviations

Above the critical scale, the bulk variables become asymptotically independent and the maximum satisfies a law of large numbers and a central limit theorem. That is,

## Proposition

If  $N - \rho_c L \gg \sqrt{L \log L}$ , then

$$\frac{M_L}{N - \rho_c L} \xrightarrow{\mu_{L,N}} 1, \quad \text{and}$$

$$\frac{M_L - (N - \rho_c L)}{\sqrt{\sigma^2 L}} \xrightarrow{d} \mathcal{N}(0, 1).$$

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**Supercritical behaviour:** Typical size of the maximum is  $N - \rho_c L$  with **Gaussian** fluctuations.

# Large Deviations (downside)

Below the lower critical scale, the asymptotic distribution of the maximum under  $\mu_{L,N}$  becomes Gumbel.

## Proposition

If  $\rho_c L - N \gg L^{\frac{b-2}{b-1}}$ , then

$$\mu_{L,N} \left[ M_L \leq A_L + x B_L \right] \rightarrow e^{-e^{-x}},$$

where

$$A_L \sim (b-1)L^{\frac{1}{b-1}} \frac{\log J_L}{J_L}, \quad B_L \sim \frac{L^{\frac{1}{b-1}}}{J_L}, \quad J_L = \frac{\rho_c L - N}{L^{\frac{b-2}{b-1}}} \rightarrow \infty.$$

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**Subcritical behaviour:** The size of the maximum gradually reduces from  $L^{\frac{1}{b-1}}$  to  $\log L$  with **Gumbel** fluctuations.

# Lower Critical Scale

At the lower critical scale the asymptotic distribution of the maximum under  $\mu_{L,N}$  changes from Fréchet to Gumbel.

## Proposition

If  $\frac{\rho_c L - N}{L^{\frac{b-2}{b-1}}} \rightarrow \omega > 0$ , then

$$\mu_{L,N} \left[ M_L \leq x L^{\frac{1}{b-1}} \right] \longrightarrow \exp \left\{ -C \int_x^{+\infty} e^{-\omega t} \frac{dt}{t^b} \right\}, \forall x > 0.$$



# Critical Scale

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$$\lim \frac{N - \rho_c L}{\Delta_L} < 1 \quad \longrightarrow \quad \text{Critical picture}$$

$M_L \sim L^{\frac{1}{b-1}}$ , Fréchet fluctuations.

$$\lim \frac{N - \rho_c L}{\Delta_L} > 1 \quad \longrightarrow \quad \text{Supercritical picture}$$

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$$M_L \simeq N - \rho_c L, \text{ gaussian fluctuations.}$$

# Phase transition

If  $N - \rho_c L = (1 - \epsilon)\Delta_L$

- The conditional distribution of the maximum is asymptotically equivalent to that of the maximum of independent r.v.'s drawn from  $\nu_{\phi_c}$ .
- $M_L = o(\Delta_L)$ , and the number of particles in the bulk is  $\rho_c L + O(\Delta_L)$ .

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The condensate appears with the removal of particles from the bulk. The correlations shift from being entirely absorbed by the bulk, to being entirely absorbed by the maximum.

# At the critical point

Suppose now

$$N - \rho_c L = \Delta_L \left( 1 + \frac{b}{2(b-3)} \frac{\log \log L}{\log L} + \frac{\gamma L}{\log L} \right).$$

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## Proposition

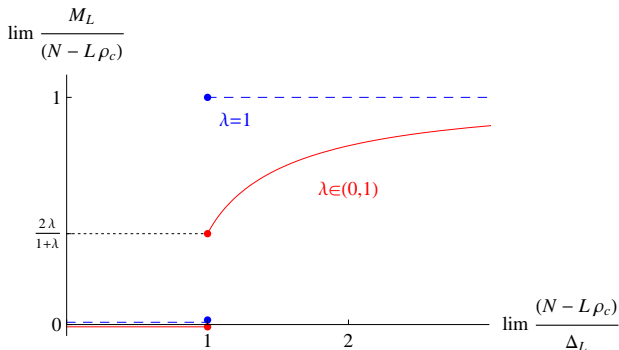
If  $\gamma_L \rightarrow \gamma \in \mathbb{R}$ ,

$$\frac{M_L}{N - \rho_c L} \xrightarrow{\mu_{L,N}} Be(p_\gamma),$$

where  $p_\gamma \in (0, 1)$  is such that  $p_\gamma \rightarrow 0(1)$  as  $\gamma \rightarrow -\infty(+\infty)$ .

# The law of large numbers when $\lambda < 1$ .

Comparison of the law of large numbers for the power law and the stretched exponential case.



# Conclusion

- We saw how large deviations for subexponential r.v.'s are typically realised: one variable assumes a big value and the bulk becomes asymptotically independent.
- We used this knowledge to obtain refined results for the typical size of the condensate and its fluctuations at supercritical densities.
- Emergence of the giant component: LLN & fluctuations for the maximum if  $N = \rho_c L + o(L)$
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**Thank you for your attention**