# On Recent Developments on Markov Processes and Applications: hand in (post lectures) task.

## My question

The Lamperti-Kiu representation for one dimensional self-similar Markov processes takes the form

$$X_t = \mathrm{e}^{\xi_{\varphi(t)}} J_{\varphi(t)}, \qquad t < \zeta$$

where  $\zeta = \inf\{t > 0 : X_t = 0\}$ ,  $\varphi(t) = \inf\{s > 0 : \int_0^s e^{\alpha \xi_u} du > t\}$  and  $(\xi, J)$  is a Markov additive Lévy process with probabilities  $\mathbf{P}_{x,i}, x \in \mathbb{R}, i \in \{-1, 1\}$ . When it exists, its matrix characteristic exponent satisfies

$$\Psi(z) = \begin{pmatrix} \psi_1(z) & 0\\ 0 & \psi_{-1}(z) \end{pmatrix} + \begin{pmatrix} -q_{1,-1} & q_{1,-1}\\ q_{-1,1} & -q_{-1,1} \end{pmatrix} \circ \begin{pmatrix} 1 & E[e^{zU_{1,-1}}]\\ E[e^{zU_{-1,1}}] & 1 \end{pmatrix},$$

where

$$\mathbf{E}_{0,i}[\mathrm{e}^{z\xi_t}; J_t = j] = \left(\mathrm{e}^{\Psi(z)t}\right)_{i,j} \qquad i, j \in \{1, -1\}$$

and  $\circ$  means element wise multiplication.

For stable processes with positive and negative jumps, it known that

$$\Psi(z) = \begin{bmatrix} \frac{\Gamma(\alpha - z)\Gamma(1 + z)}{\Gamma(\alpha\hat{\rho} - z)\Gamma(1 - \alpha\hat{\rho} + z)} & -\frac{\Gamma(\alpha - z)\Gamma(1 + z)}{\Gamma(\alpha\hat{\rho})\Gamma(1 - \alpha\hat{\rho})} \\ -\frac{\Gamma(\alpha - z)\Gamma(1 + z)}{\Gamma(\alpha\rho)\Gamma(1 - \alpha\rho)} & \frac{\Gamma(\alpha - z)\Gamma(1 + z)}{\Gamma(\alpha\rho - z)\Gamma(1 - \alpha\rho + z)} \end{bmatrix}$$

for  $\operatorname{Re}(z) \in (-1, \alpha)$ .

#### Task: Prove this identity.

## Your answer

Please present your solution fully as an essay (almost like writing a mini-paper), with a commentary on how you have developed your computations. Below are some facts that you may wish to take into consideration, in which we use the notation  $\overline{X}_t = \sup_{s \le t} X_s$  and  $\underline{X}_t = \inf_{s \le t} X_s$  for the stable process X.

• In the slides, we have considered the positive self-similar Markov process  $X_t \mathbf{1}_{(\underline{X}_t>0)}, t \ge 0$  and discussed its Lampert representation as a positive self-similar Markov process. More precisely, we showed that

$$X_t^* := X_t \mathbf{1}_{(\underline{X}_t > 0)} = e^{\xi_{\varphi^*(t)}^*}, \qquad t < \zeta^*,$$

where  $\zeta^* = \inf\{t > 0 : X_t^* = 0\}$ ,  $\varphi^*(t) = \inf\{s > 0 : \int_0^s e^{\alpha \xi_u^*} du > t\}$  and  $\xi^*$  is a Lévy process with killing at an independent and exponentially distributed random time. Moreover, we computed it characteristic exponent such that

$$E[\mathrm{e}^{\mathrm{i}\theta\xi_t^*}] = \mathrm{e}^{-\Psi^*(\theta)t}, \qquad t \ge 0$$

and

$$\Psi^*(z) = \frac{\Gamma(\alpha - iz)}{\Gamma(\alpha \hat{\rho} - iz)} \frac{\Gamma(1 + iz)}{\Gamma(1 - \alpha \hat{\rho} + iz)} = q^* + \mathsf{LK} \text{ formula of unkilled process},$$

for  $z \in \mathbb{R}$ . The killing rate  $q^* = \frac{\Gamma(\alpha)}{\Gamma(\alpha\hat{\rho})\Gamma(1-\alpha\hat{\rho})}$ .

• The distributions  $U_{-1,1}$  and  $U_{1,-1}$  can be described in terms of the path of the stable process *X*. Specifically, one should show as an intermediary step that

$$\mathbf{e}^{U_{1,-1}} =^d \frac{|X_{\tau_0^-}|}{X_{\tau_0^-}},$$

where  $\tau_0^- = \inf\{t > 0 : X_t < 0\}.$ 

• In the above claim, why does the point of issue of X not matter? This is because of the scaling property that  $(cX_{c^{-\alpha}t}, t \ge 0)$  under  $\mathbb{P}_x$  is equal in law to  $(X_t, t \ge 0)$  under  $\mathbb{P}_{cx}$ . Indeed, suppose we include the point of issue of X in its notation, writing instead  $(X_t^{(x)}, t \ge 0)$ . Then note that

$$\begin{split} \tau_0^-(x) &:= \inf\{s > 0: X_s^{(x)} < 0\} \\ &= \inf\{c^{-\alpha}(c^\alpha s) > 0: cX_{c^{-\alpha}(c^\alpha s)}^{(x)} < 0\} \\ &= c^{-\alpha}\inf\{t > 0: cX_{c^{-\alpha}t}^{(x)} < 0\} \\ &= ^d c^{-\alpha}\inf\{t > 0: X_t^{(cx)} < 0\} \\ &= c^{-\alpha}\tau_0^-(cx). \end{split}$$

As a consequence

$$X_{\tau_{0}^{-}(x)}^{(x)} = \frac{1}{c} c X_{c^{-\alpha} c^{\alpha} \tau_{0}^{-}(x)}^{(x)} =^{d} \frac{1}{c} X_{\tau_{0}^{-}(cx)}^{(cx)}$$

and hence

$$\frac{|X_{\tau_0^-(x)}^{(x)}|}{X_{\tau_0^-(x)-}} = \frac{|X_{\tau_0^-(cx)}^{(cx)}|}{X_{\tau_0^-(cx)-}^{(cx)-}}.$$

• A useful distribution to have to hand is the following: For  $y \in [0, x]$ ,  $v \ge y$  and u > 0,

$$\begin{split} \mathbb{P}(X_{\tau_x^+} - x \in \mathrm{d}u, x - X_{\tau_x^+-} \in \mathrm{d}v, x - \overline{X}_{\tau_x^+-} \in \mathrm{d}y) \\ &= \frac{\sin \alpha \rho \pi}{\pi} \frac{\Gamma(\alpha + 1)}{\Gamma(\alpha \rho) \Gamma(\alpha(1 - \rho))} \frac{(x - y)^{\alpha \rho - 1} (v - y)^{\alpha(1 - \rho) - 1}}{(v + u)^{1 + \alpha}} \,\mathrm{d}y \,\mathrm{d}v \,\mathrm{d}u. \end{split}$$

(See the quintuple law of Chapter 7 of [1], and Exercise 7.4 therein.)

## References

[1] Kyprianou A. E. (2014) *Fluctuations of Lévy processes with applications: Introductory lectures*, second edition. Universitext, Springer.