# An overview of processes with branching

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Galton-Watson processes Branching Random Walks Crump-Mode-Jagers processes Fragmentation Chains Fragmentation processes Continuous-time Galton-Watson processes Continuous-state branching processes Branching Brownian Motion Super-Brownian motion

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- A model for asexual reproduction represented by the Markov chain  $\{Z_n:n\geq 0\}$  and  $k\in\mathbb{N}$ ,
- $\blacksquare$   $Z_n$  is the number of individuals in the *n*-th generation.
- Take,  $Z_0 = k \in \mathbb{N}$ . [Usually assume k = 1].
- Thereafter, iterate from generation n to n+1 via

$$Z_{n+1} = \sum_{j=1}^{Z_n} A_j^{(n+1)},$$

where  $\{A_j^{(n+1)}\}$  are independent of  $\{Z_1, \dots, Z_n\}$  and have a common distribution  $\{p_i : i \ge 0\}$ , known as the *offspring distribution*. [Assume  $p_1 = 0$  and distribution is not defective].

What makes this a *branching* process? Momentarily incorporate the the initial value k into the notation:

$$Z_n^{(k)} = {}^d Z_n^{(1)}(1) + Z_n^{(1)}(2) + \dots + Z_n^{(1)}(k),$$

where  $Z_n^{(1)}(j)$  is an i.i.d. copy of  $Z_n^{(1)}$ . Note Markov property: for  $k, n, n' \in \mathbb{N}$ 

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• Suppose that  $q = \mathbb{P}(Z_n = 0 \text{ for some } n \in \mathbb{N})$ 

• Let  $\mathcal{F}_n = \sigma(Z_0, \cdots, Z_n)$ , then we have a martingale:

 $\mathbb{P}(Z_n=0 ext{ for some } n\in \mathbb{N}|\mathcal{F}_n)=q^{Z_n}$ 

• The constant q is a fixed point of the p.g.f. of the offspring distribution

$$q = \sum_{j=0}^{\infty} q^{i} p_{i}$$

- Conversely, any fixed point, q, of the p.g.f. of the offspring distribution makes a martingale:  $\{q^{Z_n}:n\geq 0\}$
- Note that  $q_1 = 1$  is always a root.
- Note, moreover, that

$$\left.\sum_{j=0}^{\infty} s^{i} p_{i}\right|_{s=0+} = p_{0}$$

and a little argument shows that the p.g.f. is strictly convex and hence there is a second root  $q_2 \in (0,1)$  if and only if

$$m := \sum_{i=1}^{\infty} i p_i > 1$$

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- Or m > 1, in which case  $q_2^{Z_n}$  is a uniformly integrable martingale which has a non-trivial limit with mean  $q_2 < 1$ .
- This means  $\mathbb{P}(\mathbb{Z}_n = 0 \text{ for some } n \ge 1) \in [0, 1)$  and hence, as this probability is also a fixed point, it must be equal to  $q_2$ .

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(i) If  $m \leq 1$ , (sub)critical, then  $\mathbb{P}(Z_n = 0 \text{ for some } n \geq 1) = 1$ (ii) If m > 1, supercritical, then  $\mathbb{P}(Z_n = 0 \text{ for some } n \geq 1) = q$  where  $q = \sum_{i=0}^{\infty} q^i p_i$ .

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Another martingale (with unit mean):

$$M_n := \frac{Z_n}{m^n}, \qquad n \ge 0.$$

Note

$$\mathbb{E}[M_{n+1}|\mathcal{F}_n] = \frac{1}{m^{n+1}} \mathbb{E}[\widetilde{Z}_1^{(Z_n)}|\mathcal{F}_n]$$

so it is enough to prove that

$$\mathbb{E}[Z_1^{(\ell)}] = \ell m,$$

but this is obvious.

As a positive martingale,  $M_n$  has an almost sure limit, say  $M_\infty$ . If the latter is non-trivial, then

$$Z_n^{(k)} \sim m^n M_\infty$$
 as  $n \to \infty$ .

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Theorem (Kesten–Stigum): Suppose that m > 1. The martingale  $M_n$  is  $L^1$  convergent (in particular  $\mathbb{E}(M_{\infty}) = 1$  and hence  $M_{\infty}$  is not trivial if and only

$$\sum_{i=1}^{\infty} i \log i p_i < \infty.$$

Otherwise  $M_{\infty} \equiv 0$ .

In fact, when there is  $L^1$  convergence  $\{M_{\infty} = 0\} = \{Z_n = 0 \text{ for some } n\}$ 

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• We want to build a spatial 'branching' out of the Galton–Watson process. We think of the population in generation n as random measure on  $\mathbb{R}^d$  with atomic support, each atom having unit mass: i.e. a process  $X = \{X_n(\cdot) : n \ge 0\}$ , where

$$X_n(\cdot) = \sum_{i=1}^{Z_n} \delta_{x_i^n}(\cdot),$$

and  $\{x_i^n : i = 1, \cdots, Z_n\}$  are the positions and number of particles making up the support of  $X_n$ .

- Consider a point process  $\xi(\cdot)$  on  $\mathbb{R}^d$ .
- Let  $X_0(\cdot) = \delta_0(\cdot)$  and, given  $X_n(\cdot)$ ,

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■ X is still a branching process, in the sense of independent additivity, and Markovian, as a measure-valued process.

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$$\langle 1, X_n \rangle = \int_{\mathbb{R}^d} 1 X_n(\mathrm{d}x) = Z_n$$

Pre-emptive choice of notation:  $Z_n$  is again a Galton–Watson process. Drop to one dimension, fix  $\theta \in \mathbb{R}$  and let

$$m(\theta) = \mathbb{E}\left[\sum_{i=1}^{Z_1} e^{-\theta x_i^n}\right]$$

Note that when  $\theta = 0$  then  $m(0) = m = \mathbb{E}(Z_1)$ , as before.

Another martingale:

$$W_n(\theta) := \frac{1}{m(\theta)^n} \sum_{i=1}^{Z_n} e^{-\theta x_i^n}, \qquad n \ge 0.$$

Note that when  $\theta = 0$ , then  $W_n(0) = Z_n/m^n$ .

■ Again, as a non-negative martingale, it has an almost sure limit, say  $W_{\infty}(\theta)$ .

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• Let  $\theta_1 = \inf\{\theta : m(\theta) < \infty\}$  and  $\theta_2 = \sup\{\theta : m(\theta) < \infty\}.$ 

Theorem (Biggins 1977): Suppose that  $\theta_1 < \theta_2$ . Then there exists an interval  $(\theta_*, \theta^*)$  such that, for all  $\theta \in (\theta_1, \theta_2)$ ,  $W_{\infty}(\theta)$  is an  $L^1$  limit if and only if

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otherwise  $W_{\infty}(\theta) \equiv 0$ . In fact, when there is  $L^1$  convergence

$$\{W_{\infty}(\theta)=0\}=\{Z_n=0 \text{ for some } n\}.$$

There is a remarkable connection between this theorem and the behaviour of the right most particle

$$R_n = \sup\{x_n^i : i = 1, \cdot, Z_n\} = \sup\{y \in \mathbb{R} : X_n(y, \infty) > 0\}.$$

Theorem (Biggins 1976):

$$\frac{R_n}{n} \to \gamma^* := \frac{1}{\theta^*} \log m(\theta^*) \text{ as } n \to \infty \text{ a.s on } \{\text{Extinction}\}^c$$

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# 10. Right most particle

By differentiating  $W_n(\theta)$  across its conditional expectation, one quickly establishes that

$$\partial W_n(\theta) := -\frac{\partial}{\partial \theta} W_n(\theta), \qquad n \ge 0,$$

#### is also a martingale (albeit signed).

■ Theorem (Biggins and K. 2004) (A continuation of Biggins' Martingale Convergence Theorem). For  $\theta \in [\theta_*, \theta^*]$  the derivative martingale limit exists almost surely (denoted by  $\partial W_{\infty}(\theta)$ ) and

$$\partial W_{\infty}(\theta) \equiv 0$$
 when  $\theta \in (\theta_*, \theta^*)$ ,

moreover, under some additional mild moment conditions,

 $\partial W_{\infty}(\theta_*) > 0$  and  $\partial W_{\infty}(\theta^*) > 0$ 

on  $\{Z_n = 0 \text{ for some } n\}^c$  and both have infinite mean.

Theorem (Aidekon 2012): Under mild conditions,

$$\mathbb{P}(R_n - \gamma^* n + \frac{3}{2}c^* \log n \le x) \to \mathbb{E}[\exp\{-C^* \partial W_{\infty}(\theta^*)\}]$$

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$$\infty.$$
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- Think of 'spatial displacement' as 'birth time'.
- Rather than studying the CMJ indexed by generation, it is now more natural to study the evolution the process as it evolves in 'time'.
- For example, the 'coming generation':

 $C(t) = \{ individuals born after time t whose parents were born before time t \}$ 

Denote their birth times by  $\{\sigma_u : u \in \mathcal{C}(t)\}.$ 

Malthusian Parameter: The constant  $\alpha > 0$  such that

$$\mathbb{E}\left[\int_0^\infty e^{-\alpha x} \xi(\mathrm{d}x)\right] = 1.$$

#### In fact

$$\Lambda_t(\alpha) := \sum_{u \in \mathcal{C}(t)} e^{-\alpha \sigma_u}, \qquad t \ge 0$$

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• Consider the unit interval [0, 1] fragmented randomly into smaller pieces (intervals), and the pieces arranged in descending order of their lengths:  $B_1, B_2, \cdots$ ,

$$\sum_{i\geq 1} |B_i| = 1$$

- Use independent samples from the distribution of  $(B_1, B_2, \cdots)$  to fragment each of these pieces further into further pieces. e.g. given a fragment interval I, it can be dislocated further into fragments  $(IB'_1, IB'_2, \cdots)$ , where  $(B'_1, B'_2, \cdots)$  is an independent sample from the distribution of  $(B_1, B_2, \cdots)$ .
- Suppose that  $(I_1^n, I_2^n, \cdots)$  are the pieces (intervals) in the *n*-th generation of fragmentations, arranged in decreasing order of size. Then

$$X_n(\cdot) := \sum_{j=1}^{\infty} \delta_{-\log I_j^n}(\cdot), \qquad n \ge 0,$$

is a C-M-J process.

■ Note that the process can equally be represented by a sequence of ordered length (or mass) partitions of [0, 1], indexed by generations of fragmentation:

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- A self-similar fragmentation chain has the property that a fragment of size s has an independent exponentially distributed holding time with rate which is proportional to  $s^{\alpha}$ . Here,  $\alpha \in \mathbb{R}$  is the index of self-similarity.
- $\blacksquare \ \alpha = 0$  is the homogenous case considered on the previous slide.
- In general, the resulting fragmentation chain has the property that, for any  $c\in(0,1),$

 $\{cI(c^{\alpha}t): t \geq 0\}$  with  $I(0) = (1, 0, 0, \cdots),$ 

is equal in law to

 ${I(t): t \ge 0}$  with  $I(0) = (c, 0, 0, \cdots)$ .

Branching and Markov properties still to be found: The law of I(t+s) given  $\{I(u) : u \le s\}$  is equal in law to the ordering of the collective mass partitions produced by an independent sequence of mass partitions

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- A (self-similar) mass fragmentation process is a stochastic process which is valued on the space of ordered mass partitions of [0, 1] and which satisfies the branching and Markov properties in the previous bullet point.
- A fragmentation chain fits the description of a fragmentation process, but there are more processes to be found in the latter class.
- In general, one can find fragmentation processes for which dislocation times form a dense set of of  $[0, \infty)$ .
- Fragmentation chains are to fragmentation processes what compound Poiss on processes are to Lévy processes.

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- Following the example of fragmentation chains, we can convert a Galton–Watson process to a continuous-time branching process by giving holding times (life lengths) to individuals before they branch, which are independent and identically exponentially distributed with parameter, say, β.
- Write  $\{Z(t) : t \ge 0\}$  for the number of individuals at time t.
- Temporarily introducing extra notation for the number of initial individuals:  $Z^{(k)}(t)$  satisfies  $Z^{(k)}(0) = 0$ .

We still have the branching property

$$Z^{(k)}(t) = {}^{d} Z_1^{(1)}(t) + \dots + Z_k^{(1)}(t), \qquad t \ge 0,$$

where  $Z_i^{(1)}(\cdot)$  are i.i.d. copies of  $Z^{(1)}(\cdot)$ .

The lack of memory property for each life length gives us the Markov property

$$Z(t+s) =^d \widetilde{Z}^{(Z_t)}(s), \qquad t \ge 0,$$

where  $\widetilde{Z}^{(k)}(\cdot)$  is an independent copy of  $Z^{(k)}(\cdot)$ .

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We still have the branching property

 $Z^{(k)}(t) =^{d} Z_{1}^{(1)}(t) + \dots + Z_{k}^{(1)}(t), \qquad t \ge 0,$ 

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- Lots of familiar properties when we compare to the discrete-time Galton–Watson process
- If  $m \leq 1$  then  $\mathbb{P}(Z(t) = 0$  for some t > 0) = 1
- If m > 1 then  $q := \mathbb{P}(Z(t) = 0$  for some t > 0) < 1 and  $Z(t) \to \infty$  on  $\{Z(t) = 0 \text{ for some } t > 0\}^c$ .
- $q^{Z(t)}$ ,  $t \ge 0$ , is a martingale.
- When Z(0) = 1,

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Introduce a new distribution on  $\{\pi_i : i = -1, 0, 1, 2, \cdots\}$ , where  $\pi_i = p_{i+1}$ . (The number of GW offspring minus 1).

• Write, for  $t \ge 0$ ,

$$J_t = \int_0^t Z(s) \mathrm{d}s,$$

set

$$\varphi(t) = \inf\{u > 0 : J_u > t\}$$

$$L(t) = Z(\varphi(t)), \qquad t \ge 0.$$

- Consider what happens up to the first branching time  $T_1$ :
- If Z(0) = k, then  $T_1$  is the minimum of k independent exponentially distributed random variables, each with rate q. i.e.  $T_1 \sim \exp(k\beta)$ .
- And hence,  $J_{T_1} = kT_1 \sim \exp(\beta)$ .
- Apply Markov property at time T<sub>1</sub>, when the number of individuals moves from k to k + i with probability π<sub>i</sub>, and use this same reasoning again until the second branching time.
- The time change  $Z(\varphi(t))$  has the effect of spacing out branching events with independent and identical exponentially distributed random times.
- **Said** another way:  $\{L(t) : t \ge 0\}$  is a compound Poisson process with arrival rate q and jump distribution  $F(dx) = \sum_{\substack{t=1\\ t \ge 1}}^{\infty} \pi \& (da) \ge \bullet = 0$

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The converse is also true: Suppose that  $L_t$  is a compound Poisson process with arrival rate q and jump distribution  $F(dx) = \sum_{i=-1}^{\infty} \pi \delta_i(dx)$ . Let

$$K_t = \int_0^t \frac{1}{L(s)} \mathrm{d}s, \qquad t \ge 0,$$

set

$$\theta(t) = \inf\{u > 0 : K_u > t\}$$

and define

$$Z(t) = L(\theta(t) \wedge \tau_0), \qquad t \ge 0,$$

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# 21. Continuous-state branching process (CSBP)

• A  $[0, \infty]$ -valued strong Markov process  $Z = \{Z(t) : t \ge 0\}$  with probabilities  $\{P_x : x \ge 0\}$  is called a *continuous-state branching process* if it has paths that are right-continuous with left limits and its law observes the branching property: for all  $\theta \ge 0$  and  $x, y \ge 0$ ,

$$\mathbb{E}_{x+y}(\mathrm{e}^{-\theta Z(t)}) = \mathbb{E}_x(\mathrm{e}^{-\theta Z(t)})\mathbb{E}_y(\mathrm{e}^{-\theta Z(t)}).$$

The same time change using the additive functional

$$\int_0^t Y(s) \mathrm{d}s, \qquad t \ge 0$$

makes  $Z(\varphi(t))$ ,  $t \ge 0$  a Lévy process with no negative jumps.

Similarly, given a Lévy process  $\{L(t):t\geq 0\}$  with no negative jumps, the same transform as before using the additive functional

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#### 22. CSBP semi-group

Recall that a (finite mean) Lévy process with no negative jumps is characterised through its Laplace exponent:

$$\mathbb{E}[e^{-\lambda L(t)}] = \exp\{\psi(\lambda)t\}, \qquad t \ge 0,$$

where

$$\psi(\lambda) = -a\lambda + \sigma\lambda^2 + \int_{(0,\infty)} (e^{-\lambda x} - 1 + \lambda x)\Pi(dx), \qquad \lambda \ge 0,$$

where  $a \in \mathbb{R}$ ,  $\sigma \geq 0$  and  $\Pi$  is a measure on  $(0, \infty)$  satisfying  $\int_{(0,\infty)} (1 \wedge x^2) \Pi(\mathrm{d}x) < \infty.$ 

Not easy to see the CSBP Z through a path wise construction, but some information in its semi-group: For  $\theta \ge 0$ , x > 0,

$$\mathbb{E}_x(\mathrm{e}^{-\theta Z(t)}) = \mathrm{e}^{-u_t(\theta)x},$$

where, for  $t, \theta \ge 0$ ,

$$\frac{\partial}{\partial t}u_t(\theta) + \psi(u_t(\theta)) = 0, \quad \text{and } u_0(\theta) = \theta.$$

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 For comparison, consider the semi-group of the continuous-time G–W process:

$$\mathbb{E}_x(\mathrm{e}^{-\theta Z(t)}) = v_t(\theta)x,$$

where, for  $t, \theta \geq 0$ ,

$$\frac{\partial}{\partial t}v_t(\theta) = G(v_t(\theta)), \quad \text{and } u_0(\theta) = e^{-\theta}$$

where  $G(s) = \beta \left( \sum_{j=0}^{\infty} s^j p_j - s \right).$ 

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#### As before

# $\mathbb{E}_x(Z(t)) = x e^{at}$ and $Z(t) e^{-at}$ is a martingale.

- (Sub)critical if  $a \leq 0$ . Supercritical if a > 0.
- Extinguishing:  $\{Z(t) \rightarrow 0\}$
- Extinction:  $\{Z(t) = 0 \text{ for some } t \ge 0\}$  (implies extinguishing).
- $\mathbb{P}_x(\mathsf{Extinguishing}) = \exp\{-\psi^{-1}(0)x\} \ (<1 \text{ if and only if } a > 0).$
- {Extinguishing}<sup>c</sup> = { $Z(t) \to \infty$ }.
- If {Extinction}  $\neq \emptyset$  then {Extinguishing\Extinction} =  $\emptyset$  (a.s.)
- Extinction if and only if

$$\int^{\infty} \frac{1}{\psi(\theta)} \mathrm{d}\theta < \infty$$

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- Extinguishing:  $\{Z(t) \rightarrow 0\}$
- Extinction:  $\{Z(t) = 0 \text{ for some } t \ge 0\}$  (implies extinguishing).
- $\mathbb{P}_x(\mathsf{Extinguishing}) = \exp\{-\psi^{-1}(0)x\} \ (<1 \text{ if and only if } a > 0).$
- {Extinguishing}<sup>c</sup> = { $Z(t) \to \infty$ }.
- If  $\{\text{Extinction}\} \neq \emptyset$  then  $\{\text{Extinguishing} \setminus \text{Extinction}\} = \emptyset$  (a.s.)

Extinction if and only if

$$\int^{\infty} \frac{1}{\psi(\theta)} \mathrm{d}\theta < \infty$$

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- Take a continuous-time Galton-Watson processes and make each individual execute an independent (*d*-dimensional) Brownian motion from its space-time moment of birth until branching.
- Similarly to a BRW, we can describe the process as a continuous-time atomic-valued Markov process

$$X_t(\cdot) = \sum_{i=1}^{Z(t)} \delta_{x_i(t)}(\cdot)$$

Total mass:  $\langle 1, X(t) \rangle = \int_{\mathbb{R}^d} 1X_t(\mathrm{d}x) = Z(t)$ 

Martingales:

$$e^{-\beta(m-1)t} \sum_{i=1}^{N_t} e^{-\lambda x_i(t) - \lambda^2 t/2}, \quad t \ge 0$$

Right most particle for d = 1:  $R_t := \sup\{x \in \mathbb{R} : X_t(x, \infty) > 0\}$ 

$$rac{R_t}{t} o \sqrt{2eta}$$
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# 25. Branching Brownian Motion (BBM)



- Want to construct a Markov process with values in the  $\mathcal{M}_F$ , the space of finite measures (on  $\mathbb{R}^d$ ).
- Fix  $\mu \in \mathcal{M}_F$ , fix  $n \in \mathbb{N}$ . Consider a BBM with an initial number of particles which are scattered in space according to a Poisson random field with intensity  $n\mu(\cdot)$ .
- $\blacksquare$  Normally BBM assigns unit mass to each individual. Now assign mass 1/n to each individual.
- Fix the branching rate in BBM at n.
- Impose a special offspring distribution such that the generator

$$G(s) = "\beta\left(\sum_{j=0}^{\infty} s^j p_j - s\right) " = n\left(\frac{1}{n}\psi(n(1-s))\right)$$

where

$$\psi(\lambda) = \sigma \lambda^2 + \int_{(0,\infty)} (e^{-\lambda y} - 1 + \lambda x) \nu(\mathrm{d}x)$$

Now take a 'weak' limit of the resulting BBM as n→∞ and we get a measure-valued Markov process X := {X<sub>t</sub>(·) : t ≥ 0} valued in M<sub>F</sub>. [Note X<sub>0</sub>(·) = µ(·)].

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• We can characterise the evolution of  $\mathcal{X}$  through its semi-group: For all bounded measurable  $f, t \ge 0$  and  $x \in \mathbb{R}^d$  and  $\mu \in \mathcal{M}_F$ ,

$$\mathbb{E}_{\delta_x}(\mathrm{e}^{-\langle f, \mathcal{X}_t \rangle}) =: \mathrm{e}^{-w(x,t)}$$

(branching)

$$\mathbb{E}_{\mu}(\mathrm{e}^{-\langle f, \mathcal{X}_t \rangle}) = \mathrm{e}^{-\int_{\mathbb{R}^d} w(x, t)\mu(\mathrm{d}x)}$$

$$\frac{\partial}{\partial t}w(x,t)=\frac{1}{2}\frac{\partial}{\partial x^2}w(x,t)-\psi(w(x,t)).$$

- It is straightforward to check that  $\{\langle 1, \mathcal{X}_t \rangle : t \ge 0\}$  is a CSBP. Inspection of  $\psi$  shows that it is a critical CSBP.
- An adaptation of this reasoning can produce a supercritical (subcritical) ψ-superBrownian motion.
- This construction implicitly describes how to scale a continuous-time G–W process to get a CSBP.
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