

Meromorphic Lévy Processes and a new Wiener-Hopf Monte-Carlo simulation method

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- Other motivations from queuing theory, population models etc.
- One is fundamentally interested in the joint distribution

$$P(X_t \in dx, \bar{X}_t \in dy)$$

for any Lévy process (X, P) .

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where $a \in \mathbb{R}$, $\sigma \in \mathbb{R}$ and Π is a measure concentrated on $\mathbb{R} \setminus \{0\}$ satisfying $\int_{\mathbb{R}} (1 \wedge x^2) \Pi(dx) < \infty \dots$

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- ...in which case there are fast-Fourier methods for inverting $\exp\{-\Psi(\theta)t\}$ to give $P(X_t \in dx)$.
- For the case of $P(\overline{X}_t \in dx)$, recent methods have concentrated on Fourier inversion of the, so called, Wiener-Hopf factors.

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- Recall that it turns out that one may always uniquely decompose

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where e_q is an independent and exponentially distributed random variable with rate $q > 0$ and $\underline{X}_t = \inf_{s \leq t} X_s$.

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- *If one is in possession of close analytical expressions for these factors, Fourier inversion, first in θ and then in q would be an option for accessing the law of \overline{X}_t and \underline{X}_t . However one is rarely in possession of the factors (even after 60 years of research into this topic), and even then there is the issue of the double Fourier inversion.*

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- There are no convenient formulae which contain both X_t and \bar{X}_t which could be Fourier inverted.

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- Note that

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is a Gamma (Erlang) distribution with parameters n and q and by the strong law of Large numbers, for $t > 0$,

$$g(n, n/t) = \sum_{i=1}^n \frac{t}{n} e^{(i)} \rightarrow t$$

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- Hence for a suitably large n , we have in distribution

$$(X_{\mathbf{g}(n, n/t)}, \overline{X}_{\mathbf{g}(n, n/t)}) \simeq (X_t, \overline{X}_t).$$

Indeed since t is not a jump time with probability 1, we have that $(X_{\mathbf{g}(n, n/t)}, \overline{X}_{\mathbf{g}(n, n/t)}) \rightarrow (X_t, \overline{X}_t)$ almost surely.

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- A reformulation of the Wiener-Hopf factorization states that

$$X_{\mathbf{e}_q} \stackrel{d}{=} S_q + I_q$$

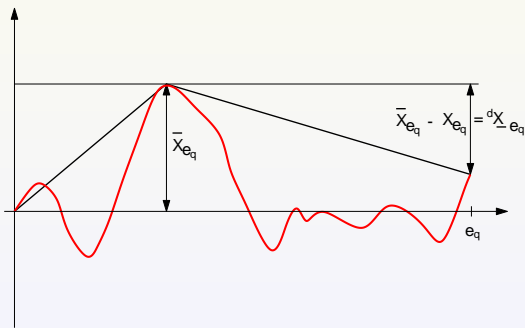
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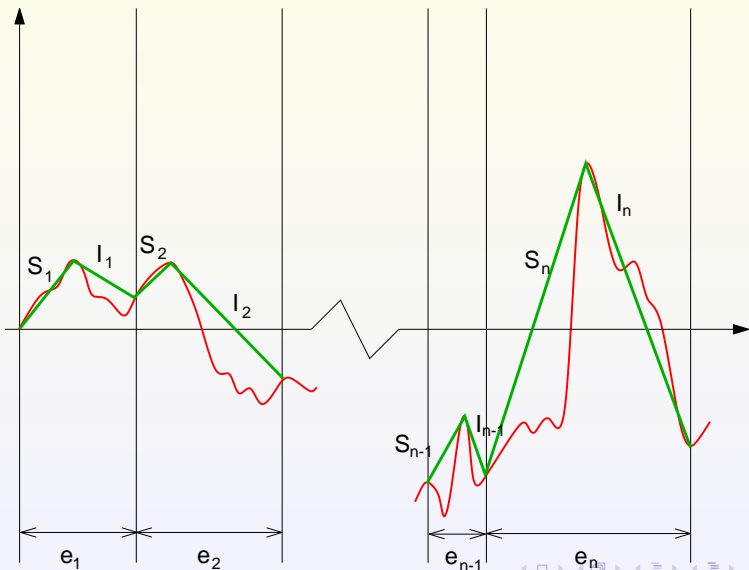
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- Taking advantage of the above, the fact that X has stationary and independent increments and the fact that, as a time, $\mathbf{g}(n, n/t)$ can be seen as n subsequent independent exponential time periods we have the following:

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- **Theorem.** For all $n \in \{1, 2, \dots\}$ and $q > 0$,

$$(X_{\mathbf{g}(n,q)}, \bar{X}_{\mathbf{g}(n,q)}) \stackrel{d}{=} (V(n, q), J(n, q))$$

where

$$V(n, q) = \sum_{j=1}^n \{S_q^{(j)} + I_q^{(j)}\} \text{ and } J(n, q) := \bigvee_{i=0}^{n-1} \left(\sum_{j=1}^i \{S_q^{(j)} + I_q^{(j)}\} + S_q^{(i+1)} \right).$$

Here, $S_q^{(0)} = I_q^{(0)} = 0$, $\{S_q^{(j)} : j \geq 1\}$ are an i.i.d. sequence of random variables with common distribution equal to that of $\bar{X}_{\mathbf{e}_q}$ and

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- **Corollary.** We have as $n \uparrow \infty$

$$(V(n, n/t), J(n, n/t)) \rightarrow (X_t, \bar{X}_t)$$

where the convergence is understood in the distributional sense.

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- This numerical procedure has disposed of one (numerical) Fourier inverse computation.
- This still leaves the problem of simulating from the unknown distribution $\bar{X}_{e_{n/t}}$ and $\underline{X}_{e_{n/t}}$ i.e. we are still one (numerical) Fourier transform away from (X_t, \bar{X}_t)

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where all the coefficients $a_n, \hat{a}_n, \rho_n, \hat{\rho}_n$ are positive, the sequences $\{\rho_n\}_{n \geq 1}$ and $\{\hat{\rho}_n\}_{n \geq 1}$ are strictly increasing, and $\rho_n \rightarrow +\infty$ and $\hat{\rho}_n \rightarrow +\infty$ as $n \rightarrow +\infty$.

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- We allow the case of a finite number summands (on either or both sides of the origin) with obvious modifications to the above.
- To ensure that $\int_{\mathbb{R}} x^2 \pi(x) dx$ converges we need to impose the additional constraint that

$$\sum_{n \geq 1} a_n \rho_n^{-2} + \sum_{n \geq 1} \hat{a}_n \hat{\rho}_n^{-2} < \infty$$

Meromorphic Lévy processes (equivalent definition)

- (i) The characteristic exponent $\Psi(z)$ is a meromorphic function which has poles at points $\{-i\rho_n, i\hat{\rho}_n\}_{n \geq 1}$, where ρ_n and $\hat{\rho}_n$ are positive real numbers.
- (ii) For $q \geq 0$ function $q + \Psi(z)$ has roots at points $\{-i\zeta_n, i\hat{\zeta}_n\}_{n \geq 1}$ where ζ_n and $\hat{\zeta}_n$ are nonnegative real numbers (strictly positive if $q > 0$). We will write $\zeta_n(q)$, $\hat{\zeta}_n(q)$ if we need to stress the dependence on q .
- (iii) The roots and poles of $q + \Psi(iz)$ satisfy the following interlacing condition

$$\dots - \rho_2 < -\zeta_2 < -\rho_1 < -\zeta_1 < 0 < \hat{\zeta}_1 < \hat{\rho}_1 < \hat{\zeta}_2 < \hat{\rho}_2 < \dots$$

- (iv) The Wiener-Hopf factors are expressed as convergent infinite products,

$$\mathbb{E} \left[e^{-z\bar{X}_{e_q}} \right] = \prod_{n \geq 1} \frac{1 + \frac{z}{\rho_n}}{1 + \frac{z}{\zeta_n}}$$

$$\mathbb{E} \left[e^{zX_{e_q}} \right] = \prod_{n \geq 1} \frac{1 + \frac{z}{\hat{\rho}_n}}{1 + \frac{z}{\hat{\zeta}_n}}.$$

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- Including Gaussian and linear drift, one can verify that the characteristic exponent is a rational function and that hyper-exponential Lévy processes have finite activity jumps and paths of bounded variation unless $\sigma > 0$.
- Note that this class has been looked at by many other authors in the past and historically it starts life as the Kou process.

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- The characteristic exponent ($\Psi(\theta) = -\log \mathbb{E}(e^{i\theta X_1}), \theta \in \mathbb{R}$) is given by

$$\begin{aligned} \Psi(\theta) = & \ iaz + \frac{1}{2}\sigma^2 z^2 + \frac{c_1}{\beta_1} \left\{ \mathbf{B}(\alpha_1, 1 - \lambda_1) - \mathbf{B}\left(\alpha_1 - \frac{i\theta}{\beta_1}, 1 - \lambda_1\right) \right\} \\ & + \frac{c_2}{\beta_2} \left\{ \mathbf{B}(\alpha_2, 1 - \lambda_2) - \mathbf{B}\left(\alpha_2 + \frac{i\theta}{\beta_2}, 1 - \lambda_2\right) \right\} \end{aligned}$$

where $\mathbf{B}(x, y) = \Gamma(x)\Gamma(y)/\Gamma(x + y)$ is the Beta function, with parameter range $a \in \mathbb{R}, \sigma, c_i, \alpha_i, \beta_i > 0$ and $\lambda_1, \lambda_2 \in (0, 3) \setminus \{1, 2\}$.

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- The corresponding Lévy measure Π has density

$$\pi(x) = c_1 \frac{e^{-\alpha_1 \beta_1 x}}{(1 - e^{-\beta_1 x})^{\lambda_1}} \mathbf{1}_{\{x>0\}} + c_2 \frac{e^{\alpha_2 \beta_2 x}}{(1 - e^{\beta_2 x})^{\lambda_2}} \mathbf{1}_{\{x<0\}}.$$

The β -class of Lévy processes includes another recently introduced family of Lévy processes known as Lamperti-stable processes.

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where $(\beta, \gamma, \hat{\beta}, \hat{\gamma})$ belong to the admissible range

$$\{\beta \leq 1, \gamma \in (0, 1), \hat{\beta} \geq 0, \hat{\gamma} \in (0, 1)\}.$$

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$$\Psi(\theta) = \frac{\Gamma(1 - \beta + \gamma - i\theta)}{\Gamma(1 - \beta + i\theta)} \frac{\Gamma(\hat{\beta} + \hat{\gamma} + i\theta)}{\Gamma(\hat{\beta} + i\theta)}$$

where $(\beta, \gamma, \hat{\beta}, \hat{\gamma})$ belong to the admissible range

$$\{\beta \leq 1, \gamma \in (0, 1), \hat{\beta} \geq 0, \hat{\gamma} \in (0, 1)\}.$$

- The Lévy density is given by

$$\pi(x) = \begin{cases} -\frac{\Gamma(\eta)}{\Gamma(\eta - \hat{\gamma})\Gamma(-\gamma)} e^{-(1-\beta+\gamma)x} {}_2F_1(1 + \gamma, \eta; \eta - \hat{\gamma}; e^{-x}) & \text{if } x > 0 \\ -\frac{\Gamma(\eta)}{\Gamma(\eta - \gamma)\Gamma(-\hat{\gamma})} e^{(\hat{\beta} + \hat{\gamma})x} {}_2F_1(1 + \hat{\gamma}, \eta; \eta - \gamma; e^{-x}) & \text{if } x < 0 \end{cases}$$

where $\eta = 1 - \beta + \gamma + \hat{\beta} + \hat{\gamma}$.

Distribution of extrema

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- For $x \geq 0$

$$\begin{aligned}\mathbb{P}(\overline{X}_{e_q} \in dx) &= \bar{a}(\rho, \zeta)^T \times \bar{v}(\zeta, x) dx \\ \mathbb{P}(-\underline{X}_{e_q} \in dx) &= \bar{a}(\hat{\rho}, \hat{\zeta})^T \times \bar{v}(\hat{\zeta}, x) dx.\end{aligned}$$

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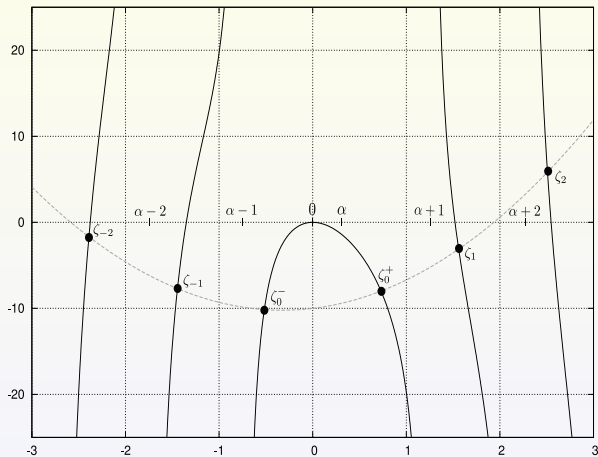
- Here $\bar{a}(\rho, \zeta) = [a_0(\rho, \zeta), a_1(\rho, \zeta), a_2(\rho, \zeta), \dots]^T$ such that

$$a_0(\rho, \zeta) = \lim_{n \rightarrow +\infty} \prod_{k=1}^n \frac{\zeta_k}{\rho_k}, \quad a_n(\rho, \zeta) = \left(1 - \frac{\zeta_n}{\rho_n}\right) \prod_{\substack{k \geq 1 \\ k \neq n}} \frac{1 - \frac{\zeta_n}{\rho_k}}{1 - \frac{\zeta_n}{\zeta_k}}$$

$$\bar{v}(\zeta, x) = \left[\delta_0(x), \zeta_1 e^{-\zeta_1 x}, \zeta_2 e^{-\zeta_2 x}, \dots \right]^T,$$

where $\delta_0(x)$ is the Dirac delta function at $x = 0$. A similar expression holds for $\bar{a}(\rho, \zeta)$.

Computing roots



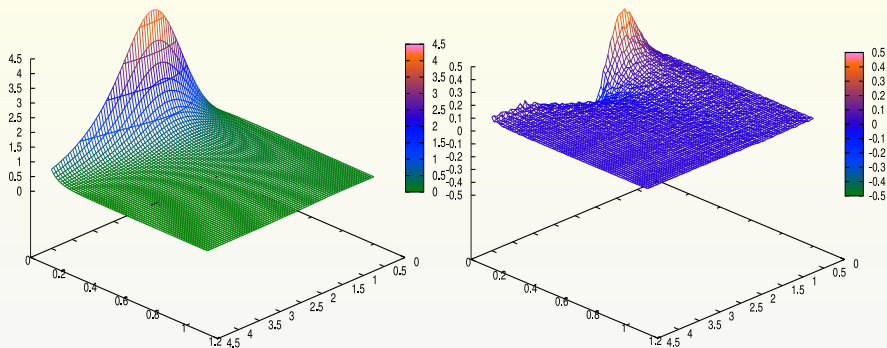


Figure: Computing the joint density of $(\bar{X}_1, \bar{X}_1 - X_1)$ for parameter Set 1. Here $\bar{X}_1 \in [0, 1]$ and $\bar{X}_1 - X_1 \in [0, 4]$. Fourier method benchmark (left), $N = 20$, WH-MC error (right).

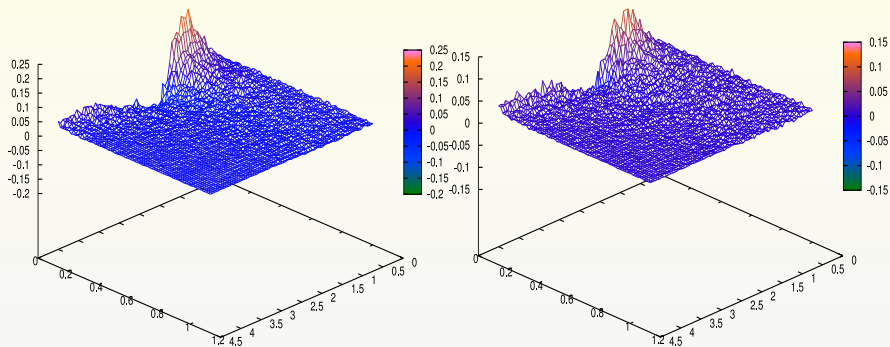


Figure: Computing the joint density of $(\bar{X}_1, \bar{X}_1 - X_1)$ for parameter Set 1. Here $\bar{X}_1 \in [0, 1]$ and $\bar{X}_1 - X_1 \in [0, 4]$. $N = 50$, WH-MC error(left), $N = 100$, WH-MC error (right).

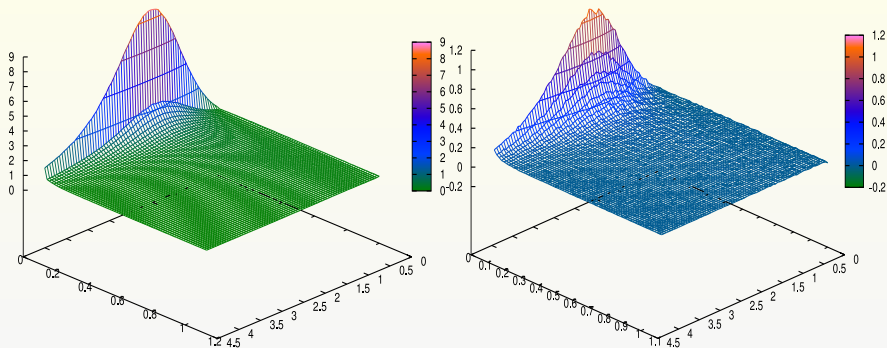


Figure: Computing the joint density of $(\bar{X}_1, \bar{X}_1 - X_1)$ for parameter Set 1. Here $\bar{X}_1 \in [0, 1]$ and $\bar{X}_1 - X_1 \in [0, 4]$. $N = 100$, MC simulation (left), $N = 100$, MC error (right).

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- Does not artificially build in an atom at zero in the numerical distribution of \bar{X}_t .

More identities: One sided exit problem

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- Define a matrix $\mathbf{A} = \{a_{i,j}\}_{i,j \geq 0}$ as

$$a_{i,j} = \begin{cases} 0 & \text{if } i = 0, j \geq 0 \\ \mathbf{a}_i(\rho, \zeta) \mathbf{b}_0(\zeta, \rho) & \text{if } i \geq 1, j = 0 \\ \frac{\mathbf{a}_i(\rho, \zeta) \mathbf{b}_j(\zeta, \rho)}{\rho_j - \zeta_i} & \text{if } i \geq 1, j \geq 1 \end{cases} \quad (2)$$

Then for $c > 0$ and $y \geq 0$ we have

$$\mathbb{E} \left[e^{-q\tau_c^+} \mathbb{I} \left(X_{\tau_c^+} - c \in dy \right) \right] = \bar{\mathbf{v}}(\zeta, c)^T \times \mathbf{A} \times \bar{\mathbf{v}}(\rho, y) dy. \quad (3)$$

More identities: Two-sided exit problem

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- Let $a > 0$ and define a matrix $\mathbf{B} = \mathbf{B}(\hat{\rho}, \zeta, a) = \{b_{i,j}\}_{i,j \geq 0}$ with

$$b_{i,j} = \begin{cases} \zeta_j e^{-a\zeta_j} & \text{if } i = 0, j \geq 1 \\ 0 & \text{if } i \geq 0, j = 0 \\ \frac{\hat{\rho}_i \zeta_j}{\hat{\rho}_i + \zeta_j} e^{-a\zeta_j} & \text{if } i \geq 1, j \geq 1 \end{cases}$$

and similarly $\hat{\mathbf{B}} = \mathbf{B}(\rho, \hat{\zeta}, a)$. There exist matrices $\mathbf{C}_1, \mathbf{C}_2$ such that for $x \in (0, a)$ we have

$$\begin{aligned} & \mathbb{E}_x \left[e^{-q\tau_a^+} \mathbb{I} \left(X_{\tau_a^+} \in dy ; \tau_a^+ < \tau_0^- \right) \right] \\ &= \left[\bar{\mathbf{v}}(\zeta, a-x)^T \times \mathbf{C}_1 + \bar{\mathbf{v}}(\hat{\zeta}, x)^T \times \mathbf{C}_2 \right] \times \bar{\mathbf{v}}(\rho, y-a) dy \end{aligned}$$

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- These matrices satisfy the following system of linear equations

$$\begin{cases} \mathbf{C}_1 &= \mathbf{A} - \hat{\mathbf{C}}_2 \mathbf{B} \mathbf{A} \\ \hat{\mathbf{C}}_2 &= -\mathbf{C}_1 \hat{\mathbf{B}} \mathbf{A} \end{cases} \quad \begin{cases} \hat{\mathbf{C}}_1 &= \hat{\mathbf{A}} - \mathbf{C}_2 \hat{\mathbf{B}} \hat{\mathbf{A}} \\ \mathbf{C}_2 &= -\hat{\mathbf{C}}_1 \mathbf{B} \hat{\mathbf{A}} \end{cases}$$

This system of linear equations can be solved iteratively with exponential convergence.

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- Define a matrix $\mathbf{D} = \{d_{i,j}\}_{i,j \geq 0}$ as follows

$$d_{i,j} = \begin{cases} 0 & \text{if } i = 0 \text{ or } j = 0 \\ \mathbf{a}_i(\rho, \zeta) \frac{\zeta_i \hat{\zeta}_j}{\zeta_i + \hat{\zeta}_j} \mathbf{a}_j(\hat{\rho}, \hat{\zeta}) & \text{if } i \geq 1, j \geq 1 \end{cases}$$

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- Then if $y \leq a$ we have

$$qR^{(q)}(a, dy) = [\bar{\mathbf{v}}(\zeta, 0 \vee y) \times \mathbf{D} \times \bar{\mathbf{v}}(\hat{\zeta}, 0 \vee (-y))] - \bar{\mathbf{v}}(\zeta, a) \times \mathbf{D} \times \bar{\mathbf{v}}(\hat{\zeta}, a - y)] dy.$$

Example of numerics

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- Choose an example from Kuznetsov's β -class that has bounded variation jump component and concentrate on four cases: With/without Gaussian component, drift to $\pm\infty$.
- For the above four cases, consider the following densities.

- (i) density of the overshoot if the exit happens at the upper boundary

$$f_1(x, y) = \frac{d}{dy} \mathbb{E}_x \left[e^{-q\tau_1^+} \mathbb{I} \left(X_{\tau_1^+} \leq y ; \tau_1^+ < \tau_0^- \right) \right]$$

- (ii) probability of exiting from the interval $[0, 1]$ at the upper boundary

$$f_2(x) = \mathbb{E}_x \left[e^{-q\tau_1^+} \mathbb{I} \left(\tau_1^+ < \tau_0^- \right) \right]$$

- (iii) probability of exiting the interval $[0, 1]$ by creeping across the upper boundary

$$f_3(x) = \mathbb{E}_x \left[e^{-q\tau_1^+} \mathbb{I} \left(X_{\tau_1^+} = 1 ; \tau_1^+ < \tau_0^- \right) \right]$$

Double sided exit: $\sigma > 0$ and positive drift

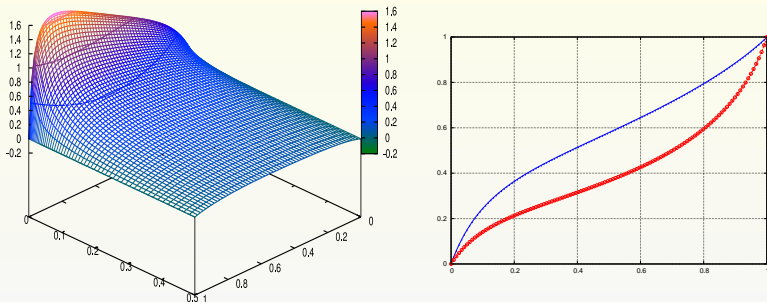


Figure: Unbounded variation case ($\sigma = 0.5$): computing the density of the overshoot $f_1(x, y)$ ($x \in (0, 1)$, $y \in (0, 0.5)$), probability of first exit $f_2(x)$ and probability of creeping $f_3(x)$ for parameter Set 1, positive drift $\mu = 1$

Double sided exit: $\sigma > 0$ and negative drift

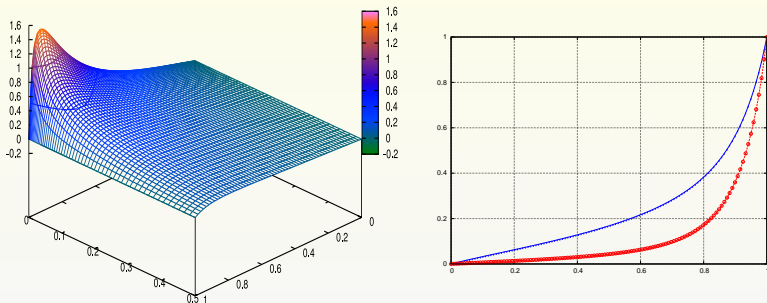


Figure: Unbounded variation case ($\sigma = 0.5$): computing the density of the overshoot $f_1(x, y)$ ($x \in (0, 1)$, $y \in (0, 0.5)$), probability of first exit $f_2(x)$ and probability of creeping $f_3(x)$ for parameter Set 2, negative drift $\mu = -1$.

Double sided exit: bounded variation and positive drift

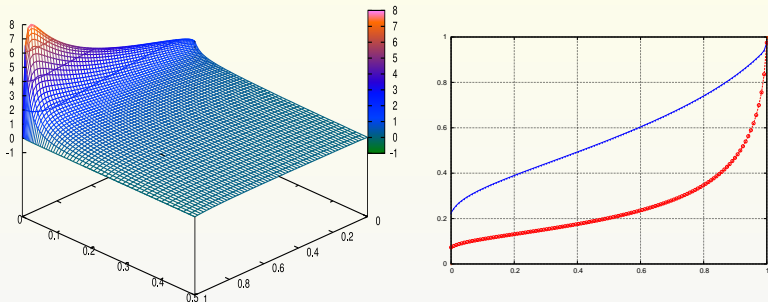


Figure: $\sigma = 0$, positive drift: computing the density of the overshoot $f_1(x, y)$ ($x \in (0, 1)$, $y \in (0, 0.5)$), probability of first exit $f_2(x)$ and probability of creeping $f_3(x)$.

Double sided exit: bounded variation and negative drift

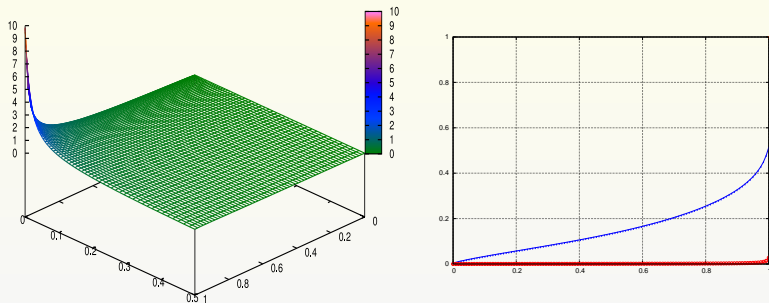


Figure: $\sigma = 0$, negative drift: computing the density of the overshoot $f_1(x, y)$ ($x \in (0, 1)$, $y \in (0, 0.5)$), probability of first exit $f_2(x)$ and probability of creeping $f_3(x)$.

Simulating processes with heavy tails

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- A little thought shows that a huge class of Lévy processes can be written as the independent sum of a β -process plus and independent compound Poisson process. Say,

$$Y_t = X_t + \sum_{i=1}^{N_t} \xi_i$$

where $\{N_t : t \geq 0\}$ is a Poisson process of rate γ and $\{\xi_i : i \geq 1\}$ and i.i.d. sequence.

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- Define iteratively for $n \geq 1$

$$V(n, \lambda) = V(n-1, \lambda) + S_{\lambda+\gamma}^{(n)} + I_{\lambda+\gamma}^{(n)} + \xi_n(1 - \beta_n)$$

$$J(n, \lambda) = \max \left(V(n, \lambda), J(n-1, \lambda), V(n-1, \lambda) + S_{\lambda+\gamma}^{(n)} \right)$$

where $V(0, \lambda) = J(0, \lambda) = 0$ and $\{\beta_n : n \geq 1\}$ are an i.i.d. sequence of Bernoulli random variables such that $\mathbb{P}(\beta_n = 1) = \lambda/(\gamma + \lambda)$. Then

$$(Y_{\mathbf{g}(n,\lambda)}, \bar{Y}_{\mathbf{g}(n,\lambda)}) \stackrel{d}{=} (V(T_n, \lambda), J(T_n, \lambda))$$

where $T_n = \min\{j \geq 1 : \sum_{i=1}^j \beta_i = n\}$.

Approximate simulation of the law of $(X_t, \overline{X}_t, \underline{X}_t)$

Define iteratively for $n \geq 1$

$$\begin{aligned} V(n, \lambda) &= V(n-1, \lambda) + S_\lambda^{(n)} + I_\lambda^{(n)} \\ J(n, \lambda) &= \max\left(J(n-1, \lambda), V(n-1, \lambda) + S_\lambda^{(n)}\right) \\ K(n, \lambda) &= \min(K(n-1, \lambda), V(n, \lambda)) \\ \tilde{J}(n, \lambda) &= \max(\tilde{J}(n-1, \lambda), V(n, \lambda)) \\ \tilde{K}(n, \lambda) &= \min\left(\tilde{K}(n-1, \lambda), V(n-1, \lambda) + I_\lambda^{(n)}\right), \end{aligned}$$

where $V(0, \lambda) = J(0, \lambda) = K(0, \lambda) = \tilde{J}(0, \lambda) = \tilde{K}(0, \lambda) = 0$. Then for any bounded function $f(x, y, z) : \mathbb{R}^3 \mapsto \mathbb{R}$ which is increasing in z -variable we have

$$\begin{aligned} \mathbb{E}[f(V(n, \lambda), J(n, \lambda), K(n, \lambda))] &\geq \mathbb{E}[f(X_{\mathbf{g}(n, \lambda)}, \overline{X}_{\mathbf{g}(n, \lambda)}, \underline{X}_{\mathbf{g}(n, \lambda)})] \\ \mathbb{E}[f(V(n, \lambda), \tilde{K}(n, \lambda), \tilde{J}(n, \lambda))] &\leq \mathbb{E}[f(X_{\mathbf{g}(n, \lambda)}, \underline{X}_{\mathbf{g}(n, \lambda)}, \overline{X}_{\mathbf{g}(n, \lambda)})]. \end{aligned}$$